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HOW DO SOCIAL SECURITY AND INCOME AFFECT THE LIVING ARRANGEMENTS OF THE ELDERLY? EVIDENCE FROM REFORMS IN MEXICO AND URUGUAY

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Abstract*

It has been shown that the social security system and other sorts of government transfers have helped poor elderly people, such as widows, to live alone in the U.S. This paper investigates whether government financial support contributed to the increase in the probability of the vulnerable elderly living alone in Latin American countries as well. Specifically, the countries that in the 1980s experienced government reforms favorable to the vulnerable elderly, Mexico and Uruguay, are examined. It is concluded that the improvement of educational attainment was mainly responsible for helping the elderly poor to live alone in rural areas in Mexico, and not the government system. On the other hand, in Uruguay, for unmarried elderly females, the increase in social security income explains most of the increase in the probability of living alone.

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1. Introduction

In the late twentieth century, there has been an increase in the number of the elderly who choose to live alone in the United States. According to previous studies, the change in the living arrangements of elderly people is enormous. For elderly widows who are considered poor, the percentage living alone increased from 18 percent in 1940 to 62 percent in 1990 (Kramarow,1995 and McGarry and Schoeni,1998). Various studies have searched for the cause of this enormous change in the family structure of the elderly, and most have concluded that having sufficient individual income is the key to explaining the increase in the percentage of the poor elderly living alone.

After Social Security was enacted in 1935 in the U.S., for instance, the program became the main source of income for elderly widows. About two-thirds of the increase in elderly widows and more than half of the increase in elderly nonmarried females living alone was caused by the increase in Social Security income after 1940 (McGarry and Schoeni, 1998 and Costa, 1999). This positive relation between these governmental supports, such as the Social Security system and other sorts of public transfers, and the percentage of widows living alone is the so-called "crowding-out" effect: as governmental financial supports increase, the percentage of the elderly living alone increases. If this effect obtains, governmental financial supports help the elderly poor to have more choices available for living arrangements and improve their quality of life.

This crowding-out effect, however, did not necessarily happen in the Latin American countries considered here. Although there were government reforms in certain countries in Latin America to help the elderly poor, and the same upward trend in the elderly living alone is observed in these countries, government supports were not necessarily responsible for producing this increase.

The countries examined in this paper, Mexico and Uruguay, undertook in the last two decades social security reforms designed to protect a certain category of the elderly. At the same time, the upward trend in the percentage of people living alone is observed for these countries.

In Mexico, about 20% of rural elderly people who are above 65 chose to live alone in 1984, whereas 24.6% chose to live alone in 1989 (Tables 1 and 2), an increase of almost 5% over a five-year period.

In Uruguay, only 25.7% of widows, 26.5% of single females, and 35.1% of divorced females who are over 65 chose to live alone in 1981. In 1995, 34.2% of widows, 30% of single females, and 40% of divorced females in the same age range chose to live alone (Tables 3 and 4). The fraction of widows who chose to live alone increased by almost 10 percentage points over these 15 years; this percentage increase is almost two times larger than what has been seen for elderly widows in the United States in the 1990s.

In Mexico, the social security reform by the Rural Solidarity Program, which started in 1973 and expanded in the late 1980s, targeted an expansion of the coverage of social security and of health care in rural areas (Mesa-Lago,1994). Health care was provided in exchange for community service of 10 days. Social security coverage clearly increased in rural areas, as shown in Tables 5 and 6. For rural males, coverage increased from 12.6% to 19.7%. For rural females, coverage jumped from 8.6% to 21.3%. The latter year's coverage is almost 2.5 times larger than the initial period's coverage.

In Uruguay, social security reform enacted in 1987 set a minimum pension (Mesa-Lago, 1994). Those helped by this reform were poor elderly, such as unmarried elderly females, who earn below the average income. The average income for these groups jumped about 30%. The improvement in social security benefits is clarified by observing the average income provided by the social security system. The average total individual income was \$1,493.91, \$1276.38, and \$1647.49 in 1981 (in 1990 US dollars) for single females, divorced females, and widows, respectively. The averages rose to \$1,926.37, \$1,670.25, and \$2,135.78 in 1995. In the same period average social security income rose from \$1,275.76, \$1,089.99, and \$1,406.91 to \$1,645.76, \$1,426.35, and \$1,823.90 for single females, divorced females, and widows (Tables 9) and 10). Both the average total individual income and the average social security income for these elderly females increased by 29%, 31%, and 30%, respectively. These increases are much more than the average income growth for the general population between 15 and 65 years old, which is 12%. Social security coverage in Uruguay did not change much over time for any subcategory of the elderly, but it has shown a slight increase. The social security system in Uruguay is one of the oldest in the Western Hemisphere and its coverage is among the highest in Latin America. The coverage for single females and widows increased from 0.907 to 0.924, by 1.9 percent, and from 0.953 to 0.974, by 2.2 percent.

What, then, has brought about this remarkable increase in the elderly living alone in rural areas of Mexico? And why has the number of elderly livings living alone increased in Uruguay? Is there any positive link between increased governmental support brought by the social security reform and the increase in the percentage of the elderly living alone, as seen in the U.S.?

The results obtained in this paper show that in Mexico, governmental financial support actually has the opposite effect on the elderly people in rural areas, even when other socioeconomic variables are taken into account. When they have social security, the elderly are inclined to live with others. In Mexico, it seems that when the elderly have more total individual income, they tend not to live alone. In Uruguay, on the other hand, when people have more individual income, they tend to live alone. In 1995, this trend is clearer than in 1981.

This paper investigates whether, in these Latin American countries, government financial support helped the vulnerable elderly to have better living conditions. I observe growth in social security income or total individual income, or coverage expansion of social security, in addition to improvement in educational attainment, as a proxy for lifetime earnings, and other demographic changes.³ The explanatory power of these variables for the increase in the percentage living alone is examined over time, using household surveys at different periods of time.

The next section summarizes the previous work on this issue, and the third section presents the model and estimation method. The fourth section briefly explains the data, and the

¹ In the Mexican data sets, reported individual income is not complete and almost half of the elderly population did not report their individual income.

² In order to see if there is any positive relation between individual income and living arrangements as observed in the U.S., individual income described by income decile is plotted with the percentage living alone for the elderly who reported their income for both countries (Graphs 1, 2, 5, and 6).

³ For the purpose of observing the effect of the demographic variables that could be related to the percentage living alone, age is graphed together with the percentage living alone (Graphs 3, 4, 7, and 8). In both countries, the percentage decreases as age increases. This is plausible because the health condition of the elderly becomes worse as they age. The relationships between living arrangements and age do not change when the patterns in the percentage of people living alone are examined by marital status (Graphs 9 and 10).

fifth section presents results and projections. The conclusion reviews the findings and and discusses their some policy implications.

2. Previous Studies

There is an extensive literature investigating the living arrangements of elderly people. In this section, the methods that have been used and the results that have been obtained are reviewed for studies that examine the determinants of living arrangements (specifically living alone vs. living with others) with a focus on income effects. There are several income effects on living arrangements of the elderly that have been considered in the past, and these effects can be categorized according to source of income: the effect through household total income, through the individual income of elderly people, through individual income of children, through private transfers, and through public transfers to elderly people.

Kotlikoff and Morris (1988) estimated the joint living choice of the elderly and their children, using as their sample the 1986 HRCA (the Hebrew Rehabilitation Center for the Aged), which covers the state of Massachusetts. They applied probit and logit models and found that both individual income of the elderly and children's income have negative effects on the probability of living with children. Other variables that show negative effects are health status, a dummy for males, a dummy for male children, and a dummy for married children. The authors conclude that the elderly tend to live alone when their own income is higher and also when their children's income is higher.

Böersch-Supan, Hajuvassiliou, Kotlikoff, and Morris (1992) used five surveys of HRCA during the period of 1982 and 1987 to see the effect of individual income of elderly people on three different choices of living arrangements: living independently (individually or with spouse), living with others, and living in an institution. They explored two of the three choices at a time using age, a dummy variable for gender, the number of children, marital status of the elderly, and various health indicators as explanatory variables. For females, they found that own income has a negative effect on living in an institution, applying a multinominal probit model.

Other researchers have investigated the effect of public transfers on living arrangements of the elderly, focusing on subcategories of the elderly.

Hoerger, Picone, and Sloan (1996) examined the effect of public subsidies on the choice of the disabled elderly for their living arrangements. They used two waves of the National Long-

Term Care Survey (NLTCS), 1984 and 1989. The multinominal probit model was applied in order to observe the effect of public subsidies on three different choices of the disabled elderly: living independently, living intergenerationally, and living in an institution. Their specification for estimation involves four different types of explanatory variables: family financial resources, such as housing wealth; factors affecting formal and informal care prices, such as price of formal care and the number of children; government subsidies and policies, such as expected health subsidy; and health and demographic variables, such as marital status and health status. They discovered a significant negative effect of expected Medicaid home health subsidies (described in dollars per week) on living with children when independent living and intergenerational living were compared. Other variables that have significant negative effects on the probability of living in extended families were nonhousing wealth and a dummy variable for being married.

The income effect on living arrangements of the elderly prior to 1940 is studied by Costa (1997). The Union Army pension program was the first major pension program in the U.S. It became universal for disabled veterans by 1910. The percentage of veterans living alone also increased during the period of between 1900 and 1910. The study linked the pension records from the 1900 and 1910 censuses to the other characteristics of men in 45 Ohio companies; this sample resembles a national sample of Union Army veterans in terms of retirement status, wealth and status as head of household. The individuals studied were veterans who were not institutionalized and had retired by 1910. Costa found that Union Army veterans tended to live alone when monthly pension income increased. Monthly pension is the only variable that shows a significant negative effect on the probability of living in extended families: among all the independent variables considered, which included age, number of children, number of children squared, poor health, former occupation, illiteracy, immigrant status, and living in the Midwest.

McGarry and Schoeni (1998) analyze the effect of Social Security and governmental income aid to elderly widows in the long run, using data compiled from the 1940, 1950, 1960, 1970, 1980, and 1990 Censuses. A multinominal logit model is applied to investigate the effect of public transfers on four different types of living arrangements: living alone, living with at least one adult child, living in an institution, and other arrangements. Social security benefits, year dummies, and dummy variables for education, such as years of schooling between 9 and 11, and years of schooling 13 or more, show significant negative effects on the probability of living with children after controlling for other explanatory variables, such as 5-year age groups, immigrant

status, life expectancy as an indicator for health status, number of children, ethnic background, and lien law. Lien laws possibly have a positive effect on the probability of living with others since children may live with their elderly parents if their inheritance is reduced by the amount of governmental support paid to their parents. McGarry and Schoeni also separately studied the effects of public transfer on two choices: living alone and living with adult children. They reached the conclusion that Social Security has a negative effect on living with others for widows in the United States.

A recent study by Costa (1999) considered living arrangements of elderly widows between 1950 and 1990 and examined the effect of Social Security benefits primarily provided by Social Security Old Age Assistance (OAA) between 1940 and 1950 on the living arrangements of elderly nonmarried females. She used as independent variables the logarithm of average yearly payment from OAA in 1940 US dollars, ethnic background, a dummy variable for low education, a year dummy, the logarithm of labor income of heads between the ages of 25 and 44, fraction of labor force that is agricultural, and immigrant status. Among all these variables, the largest significant effect on the probability of living with others, which is represented by the fraction of nonmarried elderly females living with others, was found to be the negative effect of the logarithm of the average OAA payment. Costa inferred that 40% of the actual decrease in the fraction of nonmarried elderly females living with others was caused by the increase in Social Security benefits.

Following this work, a model based on people's utility from different living arrangements is utilized in order to study the effect of social security systems in Latin American countries.

3. Model and Empirical Method

This section presents the model. Hypotheses that are considered in this paper are:

- 1. The expansion of social security explains the increase in the elderly living alone.
- 2. Changes in demographic characteristics account for increases in the elderly living alone.

These hypotheses are examined for each subcategory of the elderly because the trend of living arrangements is different for different groups. Some subgroups have shown an enormous

increase in the percentage of the elderly living alone over time. For Mexico, rural males and rural females—the groups that show the most remarkable increase in the percentage of the elderly living alone—as well as urban males and urban females are analyzed. For Uruguay, where only an urban sample is available, single females, divorced females, and widows are compared.

In order to analyze the above relationship, first, a utility function is specified in the following way:

$$U_0 = U(X_n, \varepsilon, 0)$$
, (1)

$$U_1 = U(X_p, \varepsilon, 1), \qquad (2)$$

where 1 represents living with someone, and 0 represents living alone. U is the utility function under each situation. Xp is the set of measurable personal characteristics, and ε stands for unmeasured elements of utility in each living arrangement.

People will choose to live with someone when:

$$U_1 = U(X_p, \varepsilon, 1) > U_0 = U(X_p, \varepsilon, 0)$$
 (3)

The characteristics in *Xp* include both economic and demographic factors: individual total income, dummy variables for age (the ages between 65-69, 70-74, 75-79, and over 80), gender, marital status, region (urban or rural), years of schooling (no school or some primary, primary completed, some secondary, secondary completed, some higher education), individual total income, and social security indicators (covered by social security or not, income from social security). In the estimations described below, the excluded category for age is 65-69, and for education it is "no school or some primary education." All income variables are logarithms in 1990 US dollars.

There are other variables that can affect the living arrangements of the elderly. Life expectancy has lengthened in both countries. In terms of fertility, the number of children per woman has decreased. These factors could lead to an increase in the elderly living alone. Changes in these factors, however, are too small to explain the remarkable increase in the percentage of elderly people living independently in the medium run that is analyzed here (a period of 5 to 15 years) because these variables change very slowly. Also, female labor participation has increased over time. As stated in the previous studies, daughters are more likely

than sons to live with their parents. Some studies find the daughter's labor market participation causes a decrease in the elderly living alone. However, the causal link is ambiguous. One can equally argue that because more elderly become able to live alone and value privacy, their daughters begin to work outside. Other factors, such as life expectancy and social security income, do not present the same problems for establishing causality. For example, it is hard to make an argument that the elderly became healthier or richer through social security system, because they chose to live alone. Therefore, there is no opposite causal relationship between these factors and the increase in the elderly living alone.

In addition to this view, there is another way to look at this problem. One can think that the characteristics of people living together may affect elderly people's living arrangements. Since the data set limits us to using only the variables for the elderly, the characteristics of people living with the elderly will not be considered. (Furthermore, the endogeneity of the characteristics of people living together would be another problem, if one were to include those characteristics of others as independent variables in the estimation equation.)

I use a linear model

$$U_{1i} - U_{0i} = \beta' X_{pi} + \mu_i, \qquad (4)$$

where μ is unmeasurable determinant of the utility difference between the two conditions.

The probability of living with someone becomes:

Prob
$$(U_1 > U_0) = \text{Prob } (U_1 - U_0 > 0)$$

= Prob $(\beta X_p + \mu > 0)$
= Prob $(\mu > -\beta X_p)$ (5)

When μ is normally distributed with zero mean and unit variance, the probability of living with someone is:

$$\operatorname{Prob}(\mu < \beta X_{n}) = \operatorname{F}(\beta X_{n}) \tag{6}$$

with F is the cumulative normal distribution function.

This is estimated by a probit model with 1 meaning "living with someone" and 0 meaning "living alone".

I can examine the behavior of the elderly for each year. I can also use the data sets from different years together. The regression equation then is:

$$U_{1i} - U_{0i} = \beta' X_{pi} + \gamma D + \mu_{it}$$
, (7)

where t is the time, and D is a dummy variable for year, hereafter referred to as the "year dummy". γ is expected to be negative when a dummy variable for year indicates 1 for the later year and when more elderly people have chosen to live alone recently.⁴ This year dummy captures the change in the probability of living with others if such a change occurs.

The two hypotheses are investigated using the probit model based on equation (7) with pooled data. For the first hypothesis, the probability in the initial year is calculated by setting all variables at the mean level in the initial year. Then, only the social security variable is changed to the mean value in the later year. These two estimated probabilities are compared, and the size of the change that is caused by the increase in the coverage of social security is explored. The first and second hypotheses use this method for social security income (or the social security indicator) and demographic variables, respectively. These hypotheses clarify which of the above variables had the crowding-out effect on living with others.

Mexican Data

The quality of the Mexican income variable is poor for the elderly. If the estimation is conducted with an individual income variable, almost half (45% in 1984, and 46% in 1989)⁵ of the sample must be dropped. There are two ways to tackle this problem. One way is to use an income variable available outside of the sample. For example, the average income by state,

$$U_{1i} - U_{0i} = \beta' X_{pt} + \gamma_1 D_1 + \gamma_2 D_2 + \gamma_3 D_3 + \gamma_4 D_4 + \mu_{it}$$

where $D_1,...D_4$ are cohort-specific dummies: aged 66-70, aged 71-75, aged 76-80, and aged 81+ in the initial year (Deaton, 1985) and Jappelli and Pagano (1994). In this paper, year dummies are examined.

⁴ There is another way to capture the difference across years. The variation in the behavior of each age cohort over time can be picked up from a dummy variable for each age cohort:

⁵ For Uruguay, the percentage of elderly people with missing income is 8% in 1981 and 6% in 1995.

gender, and age. Another way is to use an econometric method to correct for a possible selection problem existing in the data set. In this paper, the second method is applied.

The econometric model is:

$$I^{*}_{i} = \alpha X_{si} + \lambda_{i}$$

$$I = 1 \text{ if } I^{*}_{i} > 0$$

$$I = 0 \text{ if } I^{*}_{i} \leq 0$$

$$U_{1i} - U_{0i} = \beta X_{pi} + \mu_{i}$$

$$(9)$$

$$\Pr ob(U_{1} > U_{0} \text{ and } I_{i} > 0) = \Pr ob(\beta X_{p} + \mu > 0 \text{ and } \alpha X_{s} + \lambda > 0)$$

$$= \Phi_{2}(\alpha X_{s}, \beta X_{s}, \rho)$$

$$(10)$$

where Φ_2 is the bivariate cumulative distribution function and ρ is the covariance between μ and λ . I is an implicit function for income, and if it is greater than zero, individual income is observed. This bivariate model was presented by Poirier (1980) as an extension of the Heckman selection model (Heckman, 1979) and has been used in various settings (for example, in the investigation of the determination of union status by Abowd and Farber, 1982).

In equation (8), I=1 means the elderly person reported his/her individual income, and 0 means that individual income is missing. In equation (9), the dependent variable is the probability of living with others: 1 applies to elderly people living with others, and 0 is living alone. Only if the elderly have 1 in the first equation (8) are they included in the second equation (9). Therefore, this is a nested bivariate probit model. In order to observe the effect, at least one independent variable that explains the missing income, but is not correlated directly with living arrangements, should be included in the first equation. There are several variables that can be considered. Labor market status, such as the unemployment rate of average young people living in the same type of household or in the same region, is a possible candidate. In this paper, as explanatory variables in the first probit model, I use the average age of the members of the household, the average labor market status of the members of the household, and years of schooling of each person.

In addition, the relation between these two decisions (whether to report income and whether to live with others) is represented by the correlation between the error term in the first

equation (8) and the error term in the second equation (9). For example, if elderly people tend to live with others when they report income, λ_i and μ_i will have the same sign and the correlation between these variables should be close to one. On the other hand, if elderly people tend to live alone when they report income, λ_i and μ_i will have the opposite sign and the correlation between these variables should be close to minus one. If there is no relation between these two decisions, the correlation between these error terms should be close to zero. In fact, in the analyses of this paper, the correlation between λ_i and μ_i was very close to one for all subcategories as I will review in the next section. Therefore, elderly people in Mexico tend to live with others when they report their income. As a result, in bivariate models the coefficients of income are expected to be smaller than those in probit models without any correction for a possible selection problem.

The more details from results using a bivariate probit model and using a probit model excluding people without income are compared and examined in the fifth section.

4. Data

The data used here are household surveys collected by the Inter-American Development Bank for a larger project with which the author has been involved. For Mexico, national surveys from 1984 and 1989 are used. The individuals are persons over 65 who reported basic demographic variables. Marital status variables and social security income are not available for Mexico, but the variable which indicates the coverage of social security is available at the individual level. The household survey in 1989 contains more observations. The number of people above 65 increased from 997 in 1984 to 2,697 in 1989. In both years, about 41% of the elderly are in rural areas, the number of observations being 404 in 1984 and 1096 in 1989. Females constitute 54% of total elderly people in both years. 533 observations in 1984 and 1440 observations in 1989 are elderly females. The elderly reported most demographic and economic variables, with the exception of total individual income. In the Mexican surveys, people report their individual income by themselves, one column for each. In the regression with total individual income as one of the independent variables, less than half of the elderly females are included because only about 29% (in 1989) to 39% (in 1984) of rural females and about 35% (in 1984) to 37% (in 1989) of urban females reported positive income. 226 observations (73 in 1984 and 153 in 1989) of rural females, 590 observations (156 in 1984 and 434 in 1989) of rural males, 450

observations (116 in 1984 and 334 in 1989) of urban females, and 744 observations (207 in 1984 and 537 in 1989) of urban males are used in the regression examining the effect of total individual income on the probability of living alone.

For Uruguay, surveys from 1981 and 1995 are used; these surveys, however, cover only urban areas. For Uruguay, both marital status and income through the social security system are available. The latter year's survey contains more observations and the number of elderly increased from 3,826 to 9,806. More than half of the elderly population is female (59% in 1981 and 61% in 1995). The distribution of each marital status group of elderly females stays the same. Specifically, widows are dominant and 53% in 1981 (1,190 observations) and 52% in 1995 (3,070 observations) of elderly females are widows. In Uruguayan surveys, the majority of elderly people reported total individual income and social security income. About 92% of single females, about 85% of divorced females, and about 96% of widows reported their social security income. However, only 62% of married females in 1981, and 73% in 1995, reported individual social security income. 736 observations of single females (223 in 1981 and 513 in 1995), 376 observations of divorced females (80 in 1981 and 296 in 1995), and 4,124 observations (1,134 in 1981 and 2,990 in 1995) of widows are used in the following regression analyses.

Institutionalized elderly are not included in the following analyses.⁷

5. Results

This section explores the relation between individual incomes that are increased by governmental reform in the social security system and the percentage of the elderly living alone. Privacy is a normal good. Because of the budget constraints of elderly people, they previously could not afford to live alone or to purchase privacy. After income is increased, they may become able to live alone. If more elderly people can manage to live alone, this fact indicates better living conditions for the elderly. Government reforms targeted certain categories: the elderly people in

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⁶ It is ideal to restrict the sample to the elderly who have at least one child if there is a large variation in the number of children across elderly people, but the information on the number of children is not available.

⁷ For Mexico, household surveys do not include elderly people who are institutionalized. For the Uruguayan 1981 data, the elderly who live with more than five non-relatives are excluded, according to the definition that is used in the previous literature. There are two in 1981 who fall into this category, and here is no one in this category in 1995.

rural areas in Mexico, and elderly females (except those who were married) in Uruguay. Therefore, the positive link between increased income and the probability of living alone is expected particularly for these categories.

In Uruguay, there is a significant positive effect of the jump of social security income on the probability of unmarried elderly females' living alone. Of all elderly females, excepting those were married, this income effect (a negative effect on the percentage of elderly females living with others) is the largest for single females as opposed to divorced or widowed females. This positive relation between social security income and the probability of living alone of elderly females coincides with the prediction of previous U.S. studies. Apparently, the government helped to improve the living conditions of elderly people.

On the other hand, in Mexico, social security income is not reported as a separate variable and individual income (that should contain the information about the increase in disposable income through government aid in health care costs) is not reported by half of the elderly living in rural areas. Therefore, in order to see the relationship between governmental support and the probability of elderly people living alone, the social security indicator reported individually (0: not covered by social security, and 1: covered) is examined. Additionally, individual income is considered only for those who reported it, and the effects are examined in two ways: using only individuals who reported income, and controlled for the possible selection problem with reported income. Neither social security coverage nor individual income has the expected sign for the rural population. As described in detail below, the social security status indicates both free health care cost and poor health condition. Therefore, I need to interpret this negative effect of the social security indicator on the probability of living alone with caution. In addition, individual income after controlled for the selection problem with reporting income has a very small and insignificant effect on the probability of living alone. It appears that income does not matter much for elderly people when they make decisions on living arrangements in Mexico.8

In this section, the results from estimation are presented in the following order: first, the results from regressions using only urban population are presented for both countries and

⁸ It may be true that there has been some change in the preference of people over time. The data set used here is too short to analyze this possibility.

compared. Second, results are investigated for each subcategory (for Mexico, rural females, rural males, urban females, and urban males and for Uruguay, single females, divorced females, and widows). Finally, some outcomes from projections using the pooled regression with both years are considered and the previously stated hypotheses are investigated. All the results are reported by country.

Comparison of Regressions for Mexico and Uruguay

Before examining targeted subcategories for each country, the same specification is applied to the same subcategory—urban population by each gender—in both countries in order to compare the effect of the same explanatory variables on the probability of living with others.

In the following estimation, dependent variables are an indicator of living arrangements: 0 means living alone (with or without a spouse), and 1 means living with others. Independent variables include years of schooling (no school or some primary, primary completed, some secondary, secondary completed, some higher education); a dummy variable for age (ages between 65-69, 70-74, 75-79, and over 80); a social security indicator (0 means not covered by social security and 1 means covered by social security); a year dummy; and a constant. The excluded category for age is the age group between 65-69, and for education it is "no school or some primary education."

For Uruguay, the social security indicator is constructed in a way that is consistent with the Mexican indicator. The social security variable in this regression is a dummy variable for reported social security income: 1 means reported positive social security income, and 0 means no reported social security income. In Uruguay, about 90% of the elderly population is covered by social security (Tables 7 and 8). Therefore, the social security indicator should have about the same effect as a constant term, because most people have 1 for the social security indicator. This comparison regression includes all the urban population in the data sets. There are 1,239 urban females and 955 urban males in Mexico, and 8,204 urban females and 5,428 urban males in Uruguay used in the regression.

Tables 11-14 report the effect of the social security indicator on the probability of living with others for the urban population in both countries. The social security indicators have positive effects on the probability of living with others, with the exception of urban males in Uruguay. The same indicator has a negative but insignificant effect on the probability of living

with others for urban males in Uruguay. It is possible that urban males who are not covered and have lower income tend to live with others. Another plausible reason for this negativity is that having social security is positively correlated with age and education. In this case, since both age and education have negative effects on the probability of living with others, the coefficient of social security income can be negative even when its pure effect is positive. For urban females and males in Mexico and for urban females in Uruguay, the social security indicator does not have a significant effect on the probability of living with others, and its coefficient is positive. The year dummy has a negative and significant coefficient in urban areas in Uruguay, but not in Mexico. This implies that, as far as urban areas are considered, in Uruguay there is a significant decrease in the percentage living alone over time, but not in Mexico. Education, especially some secondary education, is an important factor in explaining the increase in the percentage of the elderly living alone in urban Mexico for both females and males. It is significant and negative in urban Uruguay as well, although the size of the effect is not as large as in urban Mexico.

Next, the results are examined for each subcategory defined above by country. First, the results using the Mexican data are reported.

Mexico

In Mexico, I am able to estimate equation (7) with data on the rural population as well, which is the category the government targeted in the reform begun in 1973. I am mainly interested in the rural population and the effect of government support on their probability of living alone. In rural areas, the people who are covered by social security are free from the health care cost. In urban areas, however, social security does not include such aid. Since the effects of the social security system on the probability of living alone for both groups, the one targeted by the government and the other which is not, are very different, these groups are analyzed separately and compared.

Overall, the coefficients on the social security indicator did not come out with the expected sign for the rural population. They are positive as shown in Tables 15-26. When rural females have social security, they tend to live with others. This positive effect is somewhat lessened in urban areas. Although the coefficients of social security are negative for both urban males and urban females when education variables are omitted from the regression, they are not significant. Therefore, there is some positive relation between education, as a proxy for lifetime earnings, and being covered by social security in urban areas.

Tables 15-26 report both coefficients and marginal effects for Mexico at the mean value of explanatory variables by sub-categorized group, rural females, rural males, urban females, and urban males. As explanatory variables, all possible demographic variables that have been used previously in the literature and that are likely to be relevant to the change in the percentage of those living alone over a five-year period are included in addition to the social security indicator. The number of children and life expectancy, which have been used in the previous studies analyzing the effect of these variables on the probability of living alone in the long run, are not likely to affect the percentage over the short run. As a result, they are not included in the regression. The dependent variable is an indicator of living arrangements as before: 0 means living alone and 1 means living with others. The explanatory variables are a dummy variable for educational attainments: no school or some primary, primary completed, some secondary, secondary completed, some higher education; a dummy variable for age: the ages between 65-69, 70-74, 75-79, and over 80; the social security indicator: 0 means not covered by social security and 1 means covered by social security; a year dummy; and a constant term. Excluded categories are no school or some primary education and the youngest age group, between 65 and 69.

In these pooled regressions, the year dummy that captures the trend in living arrangements has different outcomes for each of the subcategories. For rural females, in any regression, this year dummy comes out significant with a negative sign. This means that independent from the effects caused by other explanatory variables, there is a significant increase in the elderly living alone over this 5-year period. This trend is also shown for rural males. For females in urban areas, there is an increase in the elderly living alone again, though not for urban males. The coefficients of year dummies are smaller in urban areas than in rural areas. It is concluded that the significant trend in the fraction of the elderly living alone is found for the rural population but not for the urban population.

Among the effects of regressors on the probability of living with others, I am specifically interested in those of social security and of income. Therefore, the effects of these two variables are examined first.

Social security has a positive effect on living with others in rural areas, both for females and males, but this effect is not significant for males. This positive effect of social security is significant, though, for rural females. The estimation is conducted with and without education

variables; for rural females, the effect decreases when education variables are not included in the regression, and for urban females and males the effects decrease dramatically. Considering the fact that education is negatively correlated with the probability of living with others, social security and education are considered to be negatively correlated. This seems plausible for rural females, since covered people in rural areas tend to be poor. In urban areas, it may be true that people do not care about social security benefits when they are more educated and are likely to be wealthy. For rural males, the opposite is true and social security is positively correlated with education. The correlation between social security and income is interesting to observe and is investigated next.

The effects of social security go up when total individual income is included as a regressor for rural females and for the urban population. This indicates that social security is positively correlated with income while income is positively related with the probability of living with others. This is true for rural females and urban females. For urban males, however, income is negatively correlated with the probability of living with others and consequently social security is negatively correlated with income. Therefore, wealthy urban males are likely to be educated and also not likely to be covered by social security. This seems true when wealthy people do not have to depend on social security after retirement. For females, this scenario is not applicable, and they tend to depend on social security even when they are wealthy.

The next important factor for elderly people in deciding their living arrangements would be total individual income. Without correction for the possible selection problem with missing values of individual income, the reported individual income has a positive but insignificant effect in rural areas overall. When rural elderly have more income, they tend to live with others. Income variables have a negative effect for urban males, although this effect is not significant. This negative sign seems more plausible and consistent with the results that are obtained using the U.S. data after 1960.

For the purpose of observing the income effect more clearly, the regressions with and without educational attainment are compared. The effects of total individual income are weakened by exclusions of education variables for all subcategories. It is concluded that for rural females, rural males, and urban females, income is likely to be negatively correlated with their education. When urban males have more education, it is not clear what kind of effect higher education has on the probability of living with others. If education makes individuals more likely

to to live alone, income is positively related with education. These income effects do not, however, take into account the possible selection problem with reported income that may exist in the data set. Therefore, these effects are likely to be biased, and the interpretation of the correlation between education and income is only for the selected individuals people if more educated people tend to report their income.

The effects of other independent variables are now reviewed. In rural areas, for males, younger age came out significant with a negative sign. This is consistent with the previous interpretation of age in the literature as an indicator of health condition. When age increases, health status worsens and a spouse tends not to survive. Therefore, it is harder to live alone.

The effect of the education variable, the primary completed education, has a significant negative effect on the probability of living with others for rural females. When elderly people complete primary education, the probability of living with others decreases by 13.8 percentage points for rural males and by 10.7 percentage points for rural females.

These results from the pooled regression are now used in order to see which variable contributes the most to the increase in the probability of living alone (or the decrease in the probability of living with others). All variables are set to the mean values in the initial year, 1984, and one variable at a time is changed to its mean value in 1989. The results are presented in Table 27. For rural females, when social security coverage is changed to the mean value in the later year, the probability of living with others goes up by 0.7 percentage points, from the estimated probability of 82.7% to 83.4%. When age variables are changed with the mean value of the age dummy for the ages between 75 and 79, the probability decreases by 0.1 percentage points. For the rural population, the variable that shows the largest contribution is education. When the mean value in 1989 is used for the education dummy for primary completed education, the probability of living with others drops by 3.4 percentage points from 0.774 to 0.740 for rural males. The estimated probability in 1989 for rural males is 0.747. Therefore, the second hypothesis is not rejected and the increase in the primary completion rate brings most of the increase in the elderly living alone for rural males. However, the first hypothesis is rejected for both rural females and rural males.

In short, in rural areas in Mexico, social security has a positive but not significant effect on the probability of living with others. The effect of social security in rural areas, however, needs to be interpreted with care. The social solidarity program, which targeted the rural population, had the purpose of expanding social security coverage to the poor and to those likely to be unhealthy; this system is similar to Medicare in the U.S. The areas that this program targeted are selected based on 19 indicators of health condition and poverty. As a result, those areas had the highest rates of death by contagious and parasitic diseases and did not have water and waste disposal systems with certain qualities. Therefore, being covered by social security means that a person is likely to be poor and unhealthy. If this dummy indicates mainly poor health condition, individuals should have a tendency to live with others when they are covered by social security, and the positive effect of social security on the probability of living with others is expected. If social security frees individual income for health care, it should allow people to have more choices in living arrangements, including living alone. This predicts a negative relation between being covered by social security and living with others. With the data sets that are available, there is no way to separate out these two conflicting effects of social security on the probability of living with others. Therefore, it is possible that the insignificant effect of social security for rural males is caused by these two effects working against each other.

For rural females, some secondary education has a significant and negative effect, and for rural males, completed primary education has a significant and negative effect. Contrary to prior perceptions based on the results from the U.S. data sets, it is concluded that social security did not crowd out inter-generational living arrangements in rural areas. For urban areas, social security has a negative effect in the estimated equation without education. Crowding out may have happened in urban areas, but this effect is not significant. The small crowding-out effect is based on the fact that the social security coverage went up in urban areas with only a small increase in the percentage of elderly people living alone.

Next, the results from the biprobit model that corrects for the possible selection problem with income are considered. The biprobit model has two probit models with two discrete dependent variables: reported individual income or not, and living alone or not. The first probit model with a variable denoting reported income or not has several independent variables: the average age of the members of the household, the average labor market status of the members of the household, and individual years of schooling. The average labor market status is not related to the choice of living arrangements directly and is related to the labor market status in the region where each household resides. In this selection probit model, the dependent variable is an indicator for reporting income: 0 means income is not reported and 1 means reported. Except for

rural females, education is positively related to the probability of reported income. Therefore, when people have more education, they tend to report their income and vice versa.

After the possible selection problem with income is corrected in this way, the coefficients of individual income become smaller but still positive and insignificant for rural population and urban females. For urban males, the coefficient is negative but again smaller than that before correction and not significant. These decreases in the size of coefficients are possibly caused by the positive correlation between reported individual income and the probability of living with others. The correlation between the error term in the first probit model and the error term in the second probit model is very close to one. Therefore, elderly people in Mexico tend to live with others when they report income and this tendency contributes to the decrease in the coefficients of income. The effect of social security is still positive and insignificant for all subgroups and the size of the effects is smaller compared to the regression before correction. This decrease in the size is possibly explained by the negative correlation between living with others and having social security.

Uruguay

For Uruguay, the groups most affected by government reform are single females, divorced females, and widows. In Uruguay, information on income purely from the social security system is available for each elderly person. Therefore, it is possible to directly measure the effect of the increase in social security income caused by the government reform on the probability of the poor elderly living alone.

Both the percentage of individuals living independently and social security income went up for the elderly poor; specifically, the increase for widows is dramatic. An increase in social security income between two years for males is 9.4%, whereas that for widows is 29.6%. Moreover, males have social security income far above the average of all elderly people; the average social security income of males is above the average social security income of the elderly by 31.5% in 1981 and 23.3% in 1995. As a consequence, males are not likely to be affected by the establishment of the minimum pension in 1987. Similarly, married people are not likely to be influenced by the minimum pension, because when they live together with a spouse, the pension as a household should not be smaller than the average. Therefore, in the following analyses, the poor elderly, specifically, single females, divorced females, and widows, are

focused on in order to observe the effect of the increase in social security income likely to be induced by the settlement of minimum pension and the probability of living with others.

As before, as explanatory variables in equation (7), all the possible demographic variables that are considered to affect the probability of living alone are included, in addition to the social security income which is available for each individual in Uruguay.

The number of children and life expectancy are not likely to affect the probability of living alone over the medium run, since these variables changed by only a negligible amount (for example, the number of children of the average woman decreased by only 0.01 over a 15-year period). The health status variable that is found in the survey has no meaningful implication because, with the exception of fewer than 10 people, the rest of the elderly reported having no problems with their health.

Independent variables that are used now with the income from social security are a dummy variable for educational attainments, a dummy variable for age, and a year dummy. Excluded categories for the dummy variables are, as in the case of Mexico, no school or some primary education and the youngest age group with an age between 65 and 69.

First, the regression is examined for all females including the interacted terms (year dummies and marital status, the amount of social security income and marital status) in order to clarify if there is a statistically significant difference in the effect of social security income by marital status. If the test leads to a conclusion that there is a significant difference in the coefficients by marital status, I need to investigate the effect of social security income on the probability of living alone for each subgroup. Then, the estimation equation with the same specification is investigated by marital status for elderly females except married ones.

The effect of social security income on the probability of living with others is negative and significant for urban females in Uruguay. The effect of SS (social security) income is on average –0.064 (Table 28). Therefore, when SS income rises by \$1,000 per year, the probability of living alone increases by 4 percentage points. When the regressions with and without education variables are compared, the effect of SS income stays the same. This result implies that the correlation between education and SS income is very small when all urban females are examined.

However, when a dummy variable for married people is included, the coefficient of SS income decreases from -0. 165 to -0.252 (Table 29). Dummy variables for younger age now

have negative coefficients and also the coefficient of the year dummy is closer to zero than that without a dummy variable for married people. Married people tend to be younger and the year dummy is negatively correlated with those persons. Considering the fact that married people do not have high social security income, the decrease in the coefficients of social security income is caused by weakened influences from age dummies and the year dummy. This dummy variable for married people can be considered as an indicator of having husbands who are alive. It is not possible to tell if divorced females have living husbands or not. This, however, is irrelevant for the probability of living alone because divorced females will not live with their ex-husbands. If husbands are alive, elderly females have a strong tendency to live alone (a category that includes living with a spouse).

When total individual income is included instead of SS income, the coefficient of total income comes out to be similar to that of social security income. The coefficients of total income are slightly larger than those of SS income. In regressions both with and without education variables, total income has an even stronger negative effect on the probability of living with others. One thing that is different from the previous result is that education now has a positive effect. This is because there is a positive correlation between total individual income and education. When total individual income increases by \$1,000, the percentage living alone increases by 4 percentage points on average for urban females in Uruguay.

When the effects of both total income and SS income are examined at the same time, the coefficients of total income are significantly negative. On the other hand, the coefficients of SS income are significantly positive. The absolute size of the coefficient of total income is larger in the regression with SS income than in the one without. Taking into account the negative relation between SS income and the probability of living with others, this shows that total income is positively related with SS income. This is true because for females more than 90% of total income is from social security (Tables A.1 and A.2). Total income and SS income are almost identical. In the following analyses, these two are examined separately.

Next, a joint test is conducted in order to observe whether there is a significant difference in the effect of SS income or year dummies by marital status. For SS income, these variables are interacted with marital status: single, widow, and married. Then, a joint test of the hypothesis of having zero coefficients for all the interacted terms is performed. The statistic is distributed as chi-squared, and the result rejects the possibility of these coefficients being zero. Similarly, a

joint test is performed for year dummies. The test again rejects these variables being zero. Therefore, the effects of these variables are different by marital status and it is better to conduct the regression analyses separately for each subgroup with different marital status.

Now the effect of SS income on the probability of living with others is explored for the poor elderly females that are most likely to be influenced by the government reform: single females, divorced females, and widows. The results for the effect of SS income are the same when urban females are examined by marital status. The results are reported in Tables 31-33. The effect of social security is negative and significant for all subcategories of poor elderly females. For example, a \$1,000 increase in annual SS income yields a 9 percent increase in the probability of living alone for single females, an 8 percent increase for divorced females, and a 7 percent increase for widows. The other variables that contribute to the increase in the percentage of living alone are year dummies. The trend in the percentage of elderly females living alone is captured by the year dummy. There is a 2 to 7% increase in the elderly living alone over this 15year period of analysis that is independent from the effect of other variables. In the regression without education categories, the absolute effect of SS income is smaller for single females and widows and larger for divorced females. Education is mostly negatively correlated with the probability of living with others for single females and divorced females and is positively correlated for widows. Thus, SS income is positively related to education for single females and is negatively related for divorced females and widows.

Total income has a similar negative effect on the probability of living with others for all elderly females except those who are married. The changes in the coefficients of total income with and without education imply a positive relation between income and education for single females and a negative relation between income and education for divorced females and widows, as in the case of social security income.

Next, a projection is made using the results from the pooled regression. The estimated probabilities of living with others are compared for each subcategory. As shown in Table 34, the increase in SS income can explain almost 100% of the increases in the percentage of single females, divorced females, and widows. SS income definitely accounts for the largest portion of the increase in the percentage of living alone that occurred over this period. Other variables explain only less than 1% of the increase in the probability of living alone. The increase in the probability of the elderly living alone cannot be explained by any single variable other than SS

income. Therefore, the first hypothesis is not rejected but the second hypothesis is rejected for urban females in Uruguay.

If in the next 15 years the income from social security is doubled, and the effect of its change on the probability of living alone stays the same, 10% more single females, 9% more divorced females, and 8% more widows will live alone. Since these females represent 41% of the total elderly population, this change will bring about a 4% increase in the total elderly who choose to live alone, all other things equal.

In sum, the crowding-out effect of SS income on living with others is found for all subcategories of urban females in Uruguay and this effect is large and significant. This result is consistent with the study that examined the effect of OAA (the Old-Age Assistance program) and SSI (the Supplemental Security Income program) on the probability of living alone for widows in the U.S. in the late twentieth century using IPUMS (Integrated Public Use Microdata Series), as observed by McGarry and Schoeni (1998). The size of the crowding-out effect produced by the increase in SS income for urban females in Uruguay is similar to its effect for widows in the U.S. during the last two decades.

6. Conclusion

This paper has investigated whether the increase in social security benefits has helped protected elderly people to live alone: rural females and rural males in Mexico, and single females, divorced females, and widows in Uruguay. The results are fairly robust. In rural areas of Mexico, social security does not have a negative effect on the probability of living with others. For rural males, primary completed education, rather than social security, has a significant and negative effect on this probability. On the other hand, in Uruguay, for all subcategories who benefited most from the government reform, social security income consistently has a significant and negative effect. Therefore, the increase in social security income helps more single females, divorced females, and widows to live alone.

It is preferable for elderly people to have as many available choices of living arrangements as possible and to be able to choose to live alone when they wish. Also the fact that more poor elderly people became able to live alone indicates the improved quality of life for the poor elderly. In Uruguay, government financial support helps the elderly to choose to live alone. In Mexico, however, government transfers are not necessarily supporting the choice of living

alone. Despite the expansion of the coverage of social security in rural areas, people prefer to live with others. Only more education seems to change these traditional living arrangements and make more elderly choose to live alone in rural Mexico.

It may be plausible that social security does not have a crowding-out effect on living with others when the social security system is not a dependable financial source. The coverage of social security was less than 50% in Mexico in the late 1980s. Another reasonable explanation would be that privacy is an inferior good until income reaches a certain level. This seems true for Mexico. Graphs 1 and 2 suggest a U-curve relationship between income distribution and the percentage of people living alone. This implies that when people have more individual income they are likely to live with others up to a certain level of income.

In Mexico, although traditional living arrangements such as living in a large family seem preferred, completed primary completed education increases the probability of living alone by 10% to 13% in rural areas. In Uruguay, on the other hand, the increase in social security income by \$1,000 per year crowds out elderly people living with others by around 10% for urban females. If the policy is formed in order to enable poor elderly people to have more choices of living arrangements and to let them choose to live alone when they prefer, in Mexico, education is the key in rural areas, rather than social security in this period of analysis, and in Uruguay, the continual increase in the social security income is essential for urban females.

Apart from the influence of social security, there is also a question of the validity of measuring living standards by living arrangements in this region. Table A.4 shows the effect of individual income (log (total individual income)) on the probability of living with others in 17 Latin American countries in the late 1990s. In Argentina, Brazil, Chile, and Uruguay, the increase in income will significantly reduce the probability of living with others. In Colombia and Mexico, by contrast, the increase in income will raise the probability of living with others. The rest of the countries demonstrate the insignificant effect of income on the probability. Although it seems that not all the elderly prefer living alone when they become relatively wealthy, there is a tendency towards valuing privacy when it is possible. The question of living standard appears to require more investigation at the country level.

In this paper, the countries with social security reform enacted in the 1980s are examined in order to see the relationship between the increase in the governmental support and living arrangements in the elderly. It would also be interesting to study several countries with different

histories of their social security systems or with different income levels and observe if the preference changes over time or the limitation of income prevents people from choosing to live alone without any change in preferences in the long run.

BIBLIOGRAPHY

Abowd, J.M. and Farber, H.S. 1982. "Job Queues and the Union Status of Workers." *Industrial and Labor Relations Review*. 35 (3): 354-367.

Bernheim, D.B., Shleifer, A., and Summers, L.H. 1985. "The Strategic Bequest Motive." *Journal of Political Economy*. 93 (6): 1045-1076.

Böersch-Supan, A. and Kotlikoff, L.J. 1989. "Household Dissolution and the Choice of Alternative Living Arrangements Among Elderly Americans." In: D.A. Wise, editor. *The Economics of Aging*. Chicago, United States: University of Chicago Press.

Böersch-Supan, A. and Hajivassiliou, V.1993. "Smooth Unbiased Multivariate Probability Simulators for Maximum Likelihood Estimation of Limited Dependent Variable Models." *Journal of Econometrics*. 58 (3): 347-368.

Böersch-Supan, A., Hajivassiliou, V., Kotlikoff, L.J. *et al.* 1992. "Health, Children, and Elderly Living Arrangements: A Multiperiod-Multinominal Project Model with Unobserved Heterogeneity and Autocorrelated Errors." In: D.A. Wise, editor. *Topics in the Economics of Aging.* Chicago, United States: University of Chicago Press.

Costa, D.L. 1999. "A House of Her Own: Old Age Assistance and the Living Arrangements of Older Nonmarried Women." *Journal of Public Economics*. 72(1): 39-59.

----. 1997. "Displacing the Family: Union Army Pensions and Elderly Living Arrangements." *Journal of Political Economy*. 105 (6): 1269-1292.

Cox, D.1987. "Motives for Private Income Transfers." *Journal of Political Economy*. 95 (3): 508-546.

Deaton, A. and Paxson, C.H. 1990. "Patterns of Aging in Thailand and Cote D'Ivoire." NBER Working Paper 3436. Cambridge, United States:, National Bureau of Economic Research.

Deaton, A. 1985. "Panel Data from Time Series of Cross Sections." *Journal of Econometrics*. 30 (1-2): 109-126.

Duryea, S. and Székely, M.1998. "Labor Markets in Latin America: A Supply-Side Story." Working Paper 374. Washington, DC, United States: Inter-American Development Bank, Research Department.

Heckman, J.J. 1979. "Sample Bias as a Specification Error." *Econometrica*. 47 (1): 153-161.

Hoerger, T.J., Picone, G.A. and Sloan, F.A. 1996. "Public Subsidies, Private Provision of Care and Living Arrangements of the Elderly." *Review of Economics and Statistics*. 78 (3): 428-440.

Inter-American Development Bank.1999. *Economic and Social Progress in Latin America*, 1998-1999 Report. Washington, DC, United States: Inter-American Development Bank.

Jappelli, T. and Pagano, M.1994. "Personal Saving in Italy." In: J.M. Poterba, editor. *International Comparisons of Household Saving*. Chicago, United States: University of Chicago Press.

Kotlikoff, L.J. and Morris, J.N. 1988. "Why Don't the Elderly Live With Their Children? A New Look." NBER Working Paper 2734. Cambridge, United States: National Bureau of Economic Research.

----. 1989. "How Much Care Do the Aged Receive from Their Children? A Bimodal Picture of Contact and Assistance." In: D.A. Wise, editor. *The Economics of Aging*. Chicago, United States: University of Chicago Press.

Kramarow, E.A. 1995. "The Elderly Who Live Alone in the United States: Historical Perspectives on Household Change." *Demography*. 32 (3): 335-352.

Laitner, J. 1997. "Intergenerational and Interhousehold Economic Links." In: M.R. Rosenzweig and O. Stark, editors. *Handbook of Population and Family Economics*. Vol. 1A. Amsterdam, The Netherlands: Elsevier.

Lam, D.A. 1983. "The Economics of Household Composition in Developing Countries." Berkeley, United States: University of California. Doctoral dissertation.

Lora, E. and Márquez, G. 1998. "The Employment Problem in Latin America: Perceptions and Stylized Facts." Working Paper 371. Washington, DC, United States: Inter-American Development Bank, Research Department.

Márquez, G. and Pagés, C.1998. "Ties that Bind: Employment Protection and Labor Market Outcomes in Latin America." Working Paper 373. Washington, DC, United States: Inter-American Development Bank, Research Department, Washington, D.C.

McFadden, D. 1974. "Conditional Logit Analysis of Qualitative Choice Behavior." In: P. Zarembka, editor. *Frontiers in Econometrics*. New York, United States: Academic Press Inc.

McGarry, K. and Schoeni, R.F. 1998. "Social Security, Economic Growth, and the Rise in Independence of Elderly Widows in the 20th Century." NBER Working Paper 6511. Cambridge, United States: National Bureau of Economic Research.

Mesa-Lago, C.1986. "Comparative Study of the Development of Social Security in Latin America." *International Social Security Review.* 37: 127-152.

----. 1994. *Changing Social Security in Latin America*. Boulder, United States and London, United Kingdom: Lynne Rienner Publishers.

----.1997. La Seguridad Social en América Latina: ¿reforma o liquidación? Lima, Peru: Fundación Friedrich Ebert.

Michael, R.T., Fuchs, V.R. and Scott, S.R. 1980. "Changes in the Propensity to Live Alone: 1950-1976." *Demography*. 17 (1): 39-56.

Poirier, D. J. 1980. "Partial Observability in Bivariate Probit Models." *Journal of Econometrics*. 12 (2): 209-217.

Ruggles, S. 1996. "Living Arrangements of the Elderly in America: 1880-1980." In: T.K. Hareven, editor. *Aging and Generational Relations Over The Life Course*. Berlin, Germany and New York, United States: Walter de Gruyter.

Ruggles, S. and Goeken, R.1992. "Race and Multigenerational Family Structure, 1900-1980." In: S.J. South. and S.E. Tolnay, editors. *The Changing American Family*. Boulder, United States: Westview Press.

U.S. Department of Commerce. 1984. "International Trends and Perspectives: Aging." Washington, DC, United States: United States Department of Commerce.

World Bank.1994. "Averting the Old Age Crisis." New York, United States: World Bank/Oxford University Press.

total			rural			urban		
e group	living alone (%)	obs.	age group	living alone (%)	obs.	age group	living alone (%)	obs.
65-69	0.193	337	65-69	0.166	118	65-69	0.207	219
70-74	0.174	298	70-74	0.216	119	70-74	0.147	179
75-79	0.235	164	75-79	0.290	74	75-79	0.191	90
30-	0.275	198	80-	0.159	93	80-	0.342	105
otal	0.212	997	Total	0.202	404	Total	0.218	593
				females			males	
			age group	living alone (%)	obs.	age group	living alone (%)	obs.
			65-69	0.167	175	65-69	0.222	162
			70-74	0.167	164	70-74	0.184	134
			75-79	0.168	78	75-79	0.312	86
			80-	0.250	116	80-	0.311	82
			Total	0.187	533	Total	0.243	464
				rural: females			rural: males	
			age group	living alone (%)	obs.	age group	living alone (%)	obs.
			65-69	0.124	55	65-69	0.203	63
			70-74	0.204	63	70-74	0.231	56
			75-79	0.273	31	75-79	0.305	43
			80-	0.156	54	80-	0.164	39
			Total	0.181	203	Total	0.225	201
				ırban: females		urban: males		
			age group	living alone (%)	obs.	age group	living alone (%)	obs.
			65-69	0.186	120	65-69	0.233	99
			70-74	0.147	101	70-74	0.147	78
			75-79	0.106	47	75-79	0.318	43
			80-	0.307	62	80-	0.392	43
			Total	0.190	330	Total	0.257	263

total			rural			urban			
age group	living alone (%)	obs.	age group	living alone (%)	obs.	age group	living alone (%)	obs.	
65-69	0.216	926	65-69	0.249	370	65-69	0.194	556	
70-74	0.252	673	70-74	0.279	284	70-74	0.233	389	
75-79	0.289	500	75-79	0.291	213	75-79	0.288	287	
80-	0.168	598	80-	0.161	229	80-	0.172	369	
Total	0.228	2697	Total	0.246	1096	Total	0.215	160 ⁻	
				females		males			
			age group	living alone (%)	obs.	age group	living alone (%)	obs	
			65-69	0.228	480	65-69	0.202	446	
			70-74	0.223	366	70-74	0.292	307	
			75-79	0.298	254	75-79	0.278	246	
			80-	0.099	340	80-	0.264	258	
			Total	0.208	1440	Total	0.251	125	
				rural: females		rural: males			
			age group	living alone (%)	obs.	age group	living alone (%)	obs	
			65-69	0.270	175	65-69	0.228	198	
			70-74	0.247	148	70-74	0.316	136	
			75-79	0.284	91	75-79	0.295	122	
			80-	0.114	117	80-	0.221	112	
			Total	0.230	531	Total	0.263	56	
				urban: females		urban: males			
			age group	living alone (%)	obs.	age group	living alone (%)	obs	
			65-69	0.203	305	65-69	0.183	251	
			70-74	0.208	218	70-74	0.273	171	
			75-79	0.305	163	75-79	0.260	124	
			80-	0.091	223	80-	0.290	146	
			Total	0.196	909	Total	0.242	692	

Table 3: Percentage of the elderly living alone: Uruguay 1981

total			females			males			
age group	living alone (%)	obs.	age group	living alone (%)	obs.	age group	living alone (%)	obs.	
65-69	0.451	1341	65-69	0.430	765	65-69	0.478	576	
70-74	0.444	1102	70-74	0.398	630	70-74	0.504	472	
75-79	0.404	706	75-79	0.353	419	75-79	0.477	287	
80-	0.291	677	80-	0.223	440	80-	0.418	237	
Total	0.412	3826	Total	0.366	2254	Total	0.477	1572	
				females: single		females: married			
			age group	living alone (%)	obs.	age group	living alone (%)	obs.	
		-	65-69	0.355	82	65-69	0.562	361	
			70-74	0.241	62	70-74	0.612	214	
			75-79	0.256	43	75-79	0.634	93	
			80-	0.169	59	80-	0.527	55	
			Total	0.264	246	Total	0.584	723	
		•	females: divorced			females: widow			
			age group	living alone (%)	obs.	age group	living alone (%)	obs.	
		-	65-69	0.362	36	65-69	0.295	285	
			70-74	0.356	28	70-74	0.291	326	
			75-79	0.272	22	75-79	0.276	261	
			80-	0.500	8	80-	0.173	318	
			Total	0.351	94	Total	0.257	1190	

Note: 1 person did not report her marital status in 1981.

Table 4: Percentage of the elderly living alone: Uruguay 1995

	total			females			males		
age group	living alone (%)	obs.	age group	living alone (%)	obs.	age group	living alone (%)	obs.	
65-69	0.479	3290	65-69	0.471	1885	65-69	0.490	1405	
70-74	0.488	2611	70-74	0.456	1551	70-74	0.534	1060	
75-79	0.514	1798	75-79	0.455	1102	75-79	0.607	696	
80-	0.418	2107	80-	0.343	1412	80-	0.571	695	
Total	0.475	9806	Total	0.434	5950	Total	0.538	3856	
			females: single			females: married			
			age group	living alone (%)	obs.	age group	living alone (%)	obs	
			65-69	0.283	159	65-69	0.580	948	
			70-74	0.245	143	70-74	0.648	572	
			75-79	0.405	106	75-79	0.673	285	
			80-	0.271	147	80-	0.706	170	
			Total	0.294	555	Total	0.624	197	
				females: divorced		females: widow			
			age group	living alone (%)	obs.	age group	living alone (%)	obs	
			65-69	0.387	150	65-69	0.374	628	
			70-74	0.423	97	70-74	0.353	739	
			75-79	0.405	57	75-79	0.373	654	
			80-	0.326	46	80-	0.296	1049	
			Total	0.392	350	Total	0.342	3070	

Table 5: Percentage of the elderly covered by social security: Mexico 1984

	total		·	rural			urban	
age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs
65-69	0.266	337	65-69	0.043	118	65-69	0.384	219
70-74	0.363	298	70-74	0.137	119	70-74	0.509	179
75-79	0.195	164	75-79	0.102	74	75-79	0.268	90
80-	0.162	198	80-	0.147	93	80-	0.170	105
Total	0.262	997	Total	0.106	404	Total	0.357	593
				females			males	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs
			65-69	0.293	175	65-69	0.236	16:
			70-74	0.363	164	70-74	0.364	13
			75-79	0.138	78	75-79	0.261	86
			80-	0.123	116	80-	0.218	82
			Total	0.253	533	Total	0.273	46
				rural: females			rural: males	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	ob
			65-69	0.020	55	65-69	0.064	63
			70-74	0.110	63	70-74	0.168	50
			75-79	0.100	31	75-79	0.104	4
			80-	0.119	54	80-	0.191	3
			Total	0.086	203	Total	0.126	20
				urban: females			urban: males	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	ob
			65-69	0.414	120	65-69	0.346	99
			70-74	0.503	101	70-74	0.520	78
			75-79	0.161	47	75-79	0.428	4
			80-	0.126	62	80-	0.233	4
			Total	0.342	330	Total	0.379	26

Table 6: Percentage of the elderly covered by social security: Mexico 1989

	total			rural			urban	
age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs
65-69	0.377	926	65-69	0.251	370	65-69	0.460	55
70-74	0.327	673	70-74	0.223	284	70-74	0.398	389
75-79	0.291	500	75-79	0.204	213	75-79	0.346	28
80-	0.258	598	80-	0.108	229	80-	0.342	36
Total	0.321	2697	Total	0.205	1096	Total	0.396	160
				females			males	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs
			65-69	0.373	480	65-69	0.381	44
			70-74	0.285	366	70-74	0.385	30
			75-79	0.280	254	75-79	0.303	24
			80-	0.223	340	80-	0.306	25
			Total	0.297	1440	Total	0.352	12
				rural: females			rural: males	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	ob
			65-69	0.266	175	65-69	0.237	19
			70-74	0.232	148	70-74	0.212	13
			75-79		91	75-79	0.175	
			13-13	0.245	91			1:
			80-	0.245 0.091	117	80-	0.129	
								11
			80-	0.091	117	80-	0.129	12 11 56
			80- Total	0.091 0.213	117	80- Total	0.129 0.197	11
			80-	0.091 0.213 urban: females	531	80-	0.129 0.197 urban: males	56
			80- Total age group	0.091 0.213 urban: females covered by social security (%)	117 531 obs. 305	80- Total age group	0.129 0.197 urban: males covered by social security (%)	1: 56 ok
			80- Total age group 65-69	0.091 0.213 urban: females covered by social security (%) 0.438	117 531 obs.	Total age group 65-69	0.129 0.197 urban: males covered by social security (%) 0.487	0t 2:
			80- Total age group 65-69 70-74	0.091 0.213 urban: females covered by social security (%) 0.438 0.317	117 531 obs. 305 218	80- Total age group 65-69 70-74	0.129 0.197 urban: males covered by social security (%) 0.487 0.525	1°

Table 7: Percentage of the elderly covered by social security: Uruguay 1981

	total			females			males	
age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
65-69	0.769	1341	65-69	0.762	765	65-69	0.778	576
70-74	0.873	1102	70-74	0.844	630	70-74	0.911	472
75-79	0.917	706	75-79	0.886	419	75-79	0.962	287
80-	0.934	677	80-	0.916	440	80-	0.966	237
Total	0.855	3826	Total	0.838	2254	Total	0.880	1572
				females: single			females: married	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
			65-69	0.915	82	65-69	0.584	361
			70-74	0.903	62	70-74	0.654	214
			75-79	0.884	43	75-79	0.678	93
			80-	0.915	59	80-	0.672	55
			Total	0.907	246	Total	0.624	723
				females: divorced			females: widow	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
			65-69	0.778	36	65-69	0.940	285
			70-74	0.893	28	70-74	0.954	326
			75-79	0.864	22	75-79	0.962	261
			80-	1.000	8	80-	0.956	318
			Total	0.851	94	Total	0.953	1190

Note: 1. 1 person did not report her marital status in 1981.

^{2.} The social security coverage is based on the reported social security income. If they do not report social security income, they are considered as being not covered. Therefore, the coverage for married females may be underestimated.

Table 8: Percentage of the elderly covered by social security: Uruguay 1995

	total			females			males	
age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
65-69	0.774	3290	65-69	0.785	1885	65-69	0.759	1405
70-74	0.893	2611	70-74	0.892	1551	70-74	0.895	1060
75-79	0.940	1798	75-79	0.931	1102	75-79	0.954	696
80-	0.965	2107	80-	0.959	1412	80-	0.978	695
Total	0.877	9806	Total	0.881	5950	Total	0.871	3856
				females: single			females: married	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
			65-69	0.873	159	65-69	0.667	948
			70-74	0.924	143	70-74	0.770	572
			75-79	0.953	106	75-79	0.799	285
			80-	0.959	147	80-	0.846	170
			Total	0.924	555	Total	0.731	1975
				females: divorced			females: widow	
			age group	covered by social security (%)	obs.	age group	covered by social security (%)	obs.
			65-69	0.752	150	65-69	0.949	628
			70-74	0.906	97	70-74	0.978	739
			75-79	0.929	57	75-79	0.985	654
			80-	0.913	46	80-	0.979	1049
			Total	0.845	350	Total	0.974	3070

Table 9: Average social security income of the elderly: Uruguay 1981

income		

	total				female	S			males		
age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.
65-69	1897.19	1962.68	1031	65-69	1439.91	1396.44	583	65-69	2491.88	2389.29	448
70-74	1807.91	1645.37	962	70-74	1393.79	1116.38	532	70-74	2319.68	2011.17	430
75-79	1712.96	1661.29	647	75-79	1391.16	1150.71	371	75-79	2145.21	2091.22	276
80-	1531.72	1335.98	632	80-	1381.57	1156.29	403	80-	1795.83	1572.27	229
Total	1763.93	1707.30	3272	Total	1404.90	1223.27	1889	Total	2253.93	2106.41	1383
					females: si	ingle			females: ma	rried	
				age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.
				65-69	1528.39	1537.64	75	65-69	1230.45	1086.65	211
				70-74	1494.88	1181.72	56	70-74	1263.83	881.93	140
				75-79	1071.65	593.74	38	75-79	1338.76	1175.66	63
				80-	1383.74	698.70	54	80-	1122.32	421.68	37
				Total	1407.00	1155.54	223	Total	1247.10	999.38	451
					females: div	orced			females: wi	dow	
				age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs
				65-69	984.05	397.10	28	65-69	1628.99	1598.70	268
				70-74	1292.03	622.09	25	70-74	1442.36	1223.11	311
				75-79	1223.77	473.53	19	75-79	1465.46	1234.87	251
				80-	2417.52	2500.66	8	80-	1385.46	1219.29	304
				Total	1280.59	969.08	80	Total	1476.33	1324.29	113

Note: 1. 1 person did not report her marital status in 1981. Her reported social security income is \$1102.48.

Table 10: Average social security income of the elderly: Uruguay 1995

annual	income	in	1000	110	dollar
annuai	income	II 1	1990	UO	uona

	total				females	3			males		
age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.
65-69	1977.10	1819.75	2549	65-69	1654.95	1541.06	1481	65-69	2424.28	2066.57	1068
70-74	2097.57	2090.81	2333	70-74	1727.15	1622.34	1384	70-74	2638.89	2534.74	949
75-79	2044.64	2053.80	1690	75-79	1711.64	1573.32	1026	75-79	2559.98	2546.59	664
80-	1905.65	1941.77	2034	80-	1697.33	1585.87	1354	80-	2321.84	2454.74	680
Total	2006.13	1972.10	8606	Total	1696.04	1580.48	5245	Total	2490.97	2383.35	3361
					females: si	ngle			females: ma	arried	
				age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.
				65-69	1795.99	2050.77	139	65-69	1355.94	1150.39	633
				70-74	1927.35	1674.73	132	70-74	1234.12	888.29	441
				75-79	1698.45	1806.40	101	75-79	1224.96	698.63	228
				80-	1684.11	1358.24	141	80-	1386.69	1229.27	144
				Total	1779.73	1732.43	513	Total	1301.25	1024.69	1446
					females: div	orced			females: w	idow	
				age group	ave. income	std.dev.	obs.	age group	ave. income	std.dev.	obs.
				65-69	1753.94	1723.10	113	65-69	1920.31	1671.63	596
				70-74	1628.80	1179.40	88	70-74	2002.55	1910.20	723
				75-79	1609.62	2202.50	53	75-79	1893.38	1658.99	644
				80-	1735.88	1453.22	42	80-	1741.02	1659.53	1027
				Total	1688.38	1639.73	296	Total	1872.82	1727.93	2990

Table 11

Probit of Probability of Living with Others: pooled 0: living alone 1: living with others

Mexico: 1984, 1989 Urban females

	MICKICO. 100	Ŧ, 1000		
	Urban femal	es	N=1,239	
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Education(primary completed)	-0.185	0.157	-0.053	0.182
Education(some secondary)	-0.545*	0.289	-0.178	0.043
Education (secondary or preparatory completed)	-0.370*	0.178	-0.115	0.034
Education (some higher education)	-0.442*	0.047	-0.141	0.013
Age (70-74)	0.074	0.092	0.020	0.290
Age (75-79)	-0.130	0.319	-0.037	0.164
Age (80+)	-0.021	0.431	-0.006	0.234
Covered by social security	0.026	0.122	0.007	0.343
Year dummy:1989	-0.018	0.044	-0.005	0.525
Constant	0.955*	0.039		

Note:1. Standard errors are adjusted for correlations within year.

- 2 The coefficients of independent variables are significant at the 5 percent level (*).
- 3.The social security coverage is based on the reported social security income. If the elderly people do not report social security income,

they are considered as being not covered. Therefore, the coverage for married females may be underestimated.

Table 12

Probit of Probability of Living with Others: pooled
0: living alone 1: living with others

Mexico: 1984, 1989

Urban males

N= 9

N= 955

	•	Standard		
Variables	Coefficients	Errors	Derivatives	Means
Education(primary completed)	-0.065	0.460	-0.021	0.166
Education(some secondary)	-0.536*	0.085	-0.191	0.047
Education (secondary or preparatory completed)	-0.224	0.455	-0.075	0.027
Education (some higher education)	-0.618	0.486	-0.223	0.049
Age (70-74)	-0.046	0.279	-0.015	0.259
Age (75-79)	-0.302*	0.011	-0.101	0.148
Age (80+)	-0.449*	0.0003	-0.151	0.232
Covered by social security	0.028	0.169	0.009	0.427
Year dummy:1989	0.089*	0.005	0.028	0.531
Constant	0.869*	0.041		

Table 13

Probit of Probability of Living with Others: pooled
0: living alone 1: living with others
Uruguay: 1981,1995
Urban females
N

	Oruguay. 190	1,1333		
	Urban female	s	N=8,204	
	5	Standard		
Variables	Coefficients E	Frrors	Derivatives	Means
Education(primary completed)	0.061*	0.023	0.024	0.322
Education(some secondary)	-0.108*	0.053	-0.042	0.088
Education (secondary or preparatory completed)	-0.159*	0.069	-0.063	0.017
Education (some higher education)	-0.163*	0.075	-0.064	0.050
Age (70-74)	0.038*	0.014	0.015	0.268
Age (75-79)	0.079	0.069	0.030	0.185
Age (80+)	0.393*	0.113	0.147	0.220
Covered by social security	0.130	0.093	0.051	0.863
Year dummy:1995	-0.190*	0.015	-0.073	0.589
Constant	0.132	0.128		

Table 14
Probit of Probability of Living with Others: pooled
0: living alone 1: living with others
Uruguay: 1981,1995

	iay: 1901,1995			
l	Jrban males		N=5,428	
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.087*	0.008	0.035	0.286
Education(some secondary)	-0.108*	0.0003	-0.043	0.099
Education (secondary or preparatory completed)	-0.042	0.032	-0.017	0.025
Education (some higher education)	-0.081	0.056	-0.032	0.060
Age (70-74)	-0.078*	0.008	-0.031	0.286
Age (75-79)	-0.143	0.130	-0.057	0.181
Age (80+)	-0.039	0.149	-0.016	0.167
Covered by social security	-0.115	0.087	-0.046	0.875
Year dummy:1995	-0.151*	0.001	-0.060	0.570
Constant	0.204	0.122		

Table 15 (Rural)Probit of Probability of Living with Others: pooled

0: living alone 1: living with others

Mexico: 1984,1989

	Rural females	3	N=734		Rural female	s	N=734	
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Total individual income (in 1990 US dollars)								
Education(primary completed)	-0.341	0.568	-0.107	0.055				
Education(some secondary)	-2.233*	1.041	-0.718	0.003				
Education (secondary or preparatory completed)	-0.32	0.314	-0.101	0.005				
Education (some higher education)								
Age (70-74)	-0.052	0.171	-0.015	0.307	-0.076	0.190	-0.022	0.307
Age (75-79)	-0.254	0.213	-0.076	0.150	-0.245	0.245	-0.074	0.150
Age (80+)	0.287	0.389	0.075	0.239	0.291	0.387	0.077	0.239
Covered by social security	0.271*	0.146	0.070	0.154	0.239*	0.076	0.063	0.154
Year dummy:1989	-0.228*	0.006	-0.063	0.533	-0.200*	0.013	-0.056	0.533
Constant	0.930*	0.154			0.895*	0.198		

	Rural females	3	N=	226
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Total individual income (in 1990 US dollars)	0.152	0.089	0.056	6.011
Education(primary completed)	-0.669	0.772	-0.261	0.049
Education(some secondary)				
Education (secondary or preparatory completed)	-0.514	1.247	-0.201	0.007
Education (some higher education)				
Age (70-74)	0.149	0.328	0.055	0.344
Age (75-79)	-0.806	0.660	-0.312	0.139
Age (80+)				
Covered by social security	0.471*	0.244	0.161	0.151
Year dummy:1989	-0.432*	0.112	-0.158	0.530
Constant	-0.276	0.681		

	Table 16 (Rura	I)		
	Probability of Living wit	h Others: po	ooled	
Mexico: 1984, 1989 0: livi	ing alone 1: living with	others		
	Rural females	s	N=	226
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Total individual income (in 1990 US dollars)	0.170	0.121	0.063	6.011
Education(primary completed) Education(some secondary)	-0.654	0.788	-0.255	0.049
Education (secondary or preparatory completed) Education (some higher education)	-0.099	1.345	-0.037	0.007
Age (70-74)	-0.008	0.578	-0.003	0.344
Age (75-79)	-0.873	0.890	-0.337	0.139
Age (80+)	-0.438	0.419	-0.169	0.175
Covered by social security				
Year dummy:1989	-0.450*	0.038	-0.165	0.530
Constant	-0.166	1.078		
	Rural females	5	N=	226
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Total individual income (in 1990 US dollars)	0.140	0.108	0.052	6.011
Education(primary completed)				
Education(some secondary)				
Education (secondary or preparatory completed)				
Education (some higher education)	-0.049	0.535	-0.018	0.344
Age (70-74)	-0.900	0.867	-0.347	0.139
Age (75-79)	-0.455	0.423	-0.175	0.175
Age (80+)				
Covered by social security				
Year dummy:1989	-0.399*	0.018	-0.146	0.530
Constant	-0.026	1.042		

·	Table 17 (Ru	ral)						
Probit of Pr	robability of Livir	ng with Oth	ers: pooled					
0: living	ng alone 1: living with others							
Mexico: 1984, 1989	: Biprobit Rural femal N=226	es	Biprobit Selection	N=734				
		Standard			Standard			
Variables	Coefficients	Errors	Variables	Coefficien				
Log (Total individual income (in 1990 US dollars))	0.021	0.039	Years of schooling	-0.005	0.020			
Education(primary completed)	-0.219	0.222	Mean age of the household	0.020*	0.003			
Education(some secondary)			Mean labor market status	0.209	0.130			
Education (secondary or preparatory completed)	-0.315	0.471	of the household					
Education (some higher education)			Constant	-1.654*	0.184			
Age (70-74)	0.058	0.084						
Age (75-79)	-0.134	0.122						
Age (80+)								
Covered by social security	0.029	0.085						
Year dummy:1989	0.407*	0.080						
Constant	-1.256*	0.247						

Table 18 (Rural) Probit of Probability of Living with Others: pooled 0: living alone 1: living with others Mexico: 1984, 1989 Rural males N=766 Rural males N=766 Standard Standard Coefficients Variables Errors Derivatives Means Coefficients Errors Derivatives Means Log(Total individual income (in 1990 US dollars)) Education(primary completed) -0.399* 0.193 -0.138 0.036 Education(some secondary) 0.185 0.167 0.054 0.015 Education (secondary or preparatory completed) -0.270 0.450 -0.092 0.003 Education (some higher education) Age (70-74) Age (75-79) -1.303 -0.205* 1.239 0.085 -0.484 -0.066 0.005 -0.193* 0.064 -0.062 0.281 0.281 -0.284* 0.038 -0.094 0.202 -0.255* 0.059 -0.084 0.202 Age (80+) 0.044 0.054 0.013 0.182 0.081 0.050 0.025 0.182 Covered by social security Year dummy:1989 0.240 -0.154* 0.225 0.478 0.066 0.165 0.470 0.071 0.165 -0.150* 0.042 -0.047 0.546 0.041 -0.048 0.546 0.071 0.833* 0.060 Constant 0.872 590 Rural males N= Standard Variables Log(Total individual income (in 1990 US dollars)) Coefficients 0.162 Derivatives Means 0.057 6.833 Errors 0.171 Education(primary completed) -0.698 0.477 -0.269 0.037 Education(some secondary) 0.196 0.475 0.066 0.013 Education (secondary or preparatory completed) Education (some higher education) -0.303 -1.328 -0.114 -0.488 0.004 0.006 0.526 0.931 Age (70-74) -0.223* 0.036 -0.080 0.316 Age (75-79) -0.640* 0.204 -0.242 0.165 -0.106 0.064 Age (80+) -0.288* 0.012 0.135 Covered by social security Year dummy:1989 0.189 0.344 0.168 -0.127* 0.027 -0.045 0.545

-0.323

Constant

	Table 19 (Rura			
	obability of Living w		ooled	
	ing alone 1: living v	with others		
Mexico:1984, 1989				
	Rural males		N=590	
	Kurarmaies	Standard	14-350	
Variables	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars))	0.175	0.192	0.062	6.833
	-0.713	0.192	-0.275	0.037
Education(primary completed) Education(some secondary)	-0.713 0.311	0.484	-0.275 0.100	0.037
Education(some secondary) Education (secondary or preparatory completed)	-0.168	0.690	-0.061	0.013
	-0.166 -1.356	0.277	-0.496	0.004
Education (some higher education)	-1.356 -0.219*	0.886	-0.496 -0.079	0.006
Age (70-74)				
Age (75-79)	-0.627*	0.174	-0.236	0.165
Age (80+)	-0.288*	0.019	-0.106	0.135
Covered by social security				
Year dummy:1989	-0.103*	0.015	-0.036	0.545
Constant	-0.399	1.282		
	Rural males		N=	590
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars))	0.150	0.165	0.053	6.833
Education(primary completed)				
Education(some secondary)				
Education (secondary or preparatory completed)				
Education (some higher education)				
Age (70-74)	-0.204*	0.044	-0.073	0.316
Age (75-79)	-0.591*	0.180	-0.223	0.165
Age (80+)	-0.232*	0.055	-0.085	0.135
Covered by social security				
Year dummy:1989	-0.108*	0.013	-0.038	0.545
Constant	-0.284	1.115		

Probit of Pro	Table 20 (Rural bability of Living		pooled					
0: living alone 1: living with others								
Mexico: 1984, 1989	Biprobit Biprobit Rural males Selection N=590							
		Standard			Standard			
Variables	Coefficients I	Errors	Variables	Coefficients	Errors			
Log(Total individual income (in 1990 US dollars))	0.060	0.037	Years of schooling	0.049*	0.023			
Education(primary completed)	0.002	0.222	Mean age of the household	0.031*	0.003			
Education(some secondary)	0.325	0.441	Mean labor market status	0.784*	0.191			
Education (secondary or preparatory completed)	-0.083	0.667	of the household					
Education (some higher education)	-0.594	0.727	Constant	-1.332*	0.204			
Age (70-74)	-0.081	0.080						
Age (75-79)	-0.280*	0.111						
Age (80+)	-0.229	0.123						
Covered by social security	0.094	0.104						
Year dummy:1989	0.452	0.074						
Constant	-0.564	0.262						

			1 (Urban)					
			Living with Oth					
	0: liv	ing alone 1	: living with otl	ners				
Mexico: 1984, 1989								
	Urban fema	les	N=1,239		Urban fema	les	N=1,239	
		Standard	,			Standard	,	
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars))								
Education(primary completed)	-0.185	0.157	-0.053	0.182				
Education(some secondary)	-0.545*	0.289	-0.178	0.043				
Education (secondary or preparatory completed)	-0.370*	0.178	-0.115	0.034				
Education (some higher education)	-0.442*	0.047	-0.141	0.013				
Age (70-74)	0.074	0.092	0.020	0.290	0.067	0.094	0.018	0.290
Age (75-79)	-0.130	0.319	-0.037	0.164	-0.114	0.309	-0.032	0.164
Age (80+)	-0.021	0.431	-0.006	0.234	0.005	0.404	0.001	0.234
Covered by social security	0.026	0.122	0.007	0.343	-0.047	0.107	-0.013	0.343
Year dummy:1989	-0.018	0.044	-0.005	0.525	-0.011	0.033	-0.003	0.525
Constant	0.955*	0.039			0.887*	0.006		
	Urban fema	les	N=450					
	0.00	Standard	11 100					
Variables	Coefficients	Errors	Derivatives	Means				
Log(Total individual income (in 1990 US dollars))	0.184*	0.064	0.061	6.844				
Education(primary completed)	-0.043	0.216	-0.014	0.233				
Education(some secondary)	-0.355	0.361	-0.128	0.040				
Education (secondary or preparatory completed)	-0.362*	0.162	-0.130	0.055				
Education (some higher education)								
Age (70-74)	0.178	0.204	0.058	0.299				
Age (75-79)	-0.014	0.782	-0.005	0.219				
Age (80+)	-0.388	0.218	-0.138	0.148				
Covered by social security	0.279	0.182	0.091	0.409				
Year dummy:1989	-0.330*	0.044	-0.110	0.454				
Constant	-0.569	0.628						

Table 22 (Urban)

Probit of Probability of Living with Others: pooled 0: living alone 1: living with others

Mexico: 1984, 1989

	Urban female	es	N=450	
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars))	0.184*	0.058	0.061	6.844
Education(primary completed)	0.052	0.107	0.017	0.233
Education(some secondary)	-0.256	0.320	-0.091	0.040
Education (secondary or preparatory completed) Education (some higher education)	-0.280	0.220	-0.100	0.055
Age (70-74)	0.180	0.215	0.058	0.299
Age (75-79)	-0.075	0.715	-0.025	0.219
Age (80+)	-0.445	0.285	-0.160	0.148
Covered by social security				
Year dummy:1989	-0.322*	0.037	-0.108	0.454
Constant	-0.474	0.523		
	Urban female	es	N=450	
		Standard		
Variables	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars)) Education(primary completed)	0.175*	0.058	0.058	6.844
Education(some secondary)				
Education (secondary or preparatory completed)				
Education (some higher education)				
Age (70-74)	0.176	0.213	0.057	0.299
Age (75-79)	-0.057	0.720	-0.019	0.219
Age (80+)	-0.439	0.246	-0.158	0.148
Covered by social security				
Year dummy:1989	-0.327*	0.055	-0.110	0.454
Constant	-0.428	0.545		

	Table 23 (Url	oan)						
Pro	bit of Probability of	Living with	Others: pooled					
0: living alone 1: living with others								
Mexico: 1984-1989	Biprobit		Biprobit					
	Urhan fema	loo	Selection					
	N=450	162	Selection	N=1.239				
	14-450			14-1,239				
		Standard			Standard			
Variables	Coefficients	Errors	Variables	Coefficients	Errors			
Log(Total individual income (in 1990 US dollars))	0.054*	0.024	Years of schooling	0.038*	0.010			
Education(primary completed)	0.120	0.087	Mean age of the household	0.021*	0.002			
Education(some secondary)	0.172	0.154	Mean labor market status	0.036	0.108			
Education (secondary or preparatory completed)	0.126	0.185	of the household					
Education (some higher education)			Constant	-1.571*	0.137			
Age (70-74)	0.020	0.072						
Age (75-79)	0.087	0.061						
Age (80+)	-0.003	0.093						
Covered by social security	0.066	0.059						
Year dummy:1989	0.478*	0.066						
Constant	-1.451*	0.173						

Mexico: 1984, 1989		bability of L	I (Urban) iving with Othe living with oth					
	Urban males		N=955		Urban males	6	N=955	
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Log(Total individual income (in 1990 US dollars))								
Education(primary completed)	-0.065	0.460	-0.021	0.166				
Education(some secondary)	-0.536*	0.085	-0.191	0.047				
Education (secondary or preparatory completed)	-0.224	0.455	-0.075	0.027				
Education (some higher education)	-0.618	0.486	-0.223	0.049				
Age (70-74)	-0.046	0.279	-0.015	0.259	-0.004	0.298	-0.001	0.259
Age (75-79)	-0.302*	0.011	-0.101	0.148	-0.252*	0.011	-0.084	0.148
Age (80+)	-0.449*	0.000	-0.151	0.232	-0.404*	0.042	-0.136	0.232
Covered by social security	0.028	0.169	0.009	0.427	-0.059	0.106	-0.019	0.427
Year dummy:1989	0.089*	0.005	0.028	0.531	0.058*	0.008	0.018	0.531
Constant	0.869*	0.041			0.813*	0.105		
	Urban males		N=744					
		Standard						
Variables	Coefficients	Errors	Derivatives	Means				
Log(Total individual income (in 1990 US dollars))	-0.119	0.074	-0.042	7.439				
Education(primary completed)	0.073	0.339	0.025	0.165				
Education(some secondary)	-0.336*	0.056	-0.125	0.061				
Education (secondary or preparatory completed)	0.055	0.400	0.019	0.035				
Education (some higher education)	-0.225	0.383	-0.083	0.065				
Age (70-74)	-0.088	0.343	-0.031	0.275				
Age (75-79)	-0.528*	0.057	-0.198	0.142				
Age (80+)	-1.028*	0.116	-0.388	0.166				
Covered by social security	0.181	0.226	0.063	0.473				
Year dummy:1989	0.052*	0.004	0.018	0.519				
Constant	1.571*	0.452						

	Table 25 (Urbai	n)				
Probit of	Probit of Probability of Living with Others: pooled					
Mexico: 1984, 1989	0: living alone 1: living with others					
	Urban males		N=744			
Variables	Coefficients	Standard Errors	Derivatives	Means		
Log(Total individual income (in 1990 US dollars))	-0.112*	0.055	-0.039	7.439		
Education(primary completed)	0.088	0.342	0.031	0.165		
Education(some secondary)	-0.276*	0.049	-0.102	0.061		
Education (secondary or preparatory completed)	0.100	0.444	0.034	0.035		
Education (some higher education)	-0.175	0.324	-0.064	0.065		
Age (70-74)	-0.068	0.374	-0.024	0.275		
Age (75-79)	-0.517*	0.022	-0.194	0.142		
Age (80+)	-1.029*	0.110	-0.389	0.166		
Covered by social security						
Year dummy:1989	0.064*	0.018	0.023	0.519		
Constant	1.578*	0.387				
	Urban males		N=744			
		Standard				
Variables	Coefficients	Errors	Derivatives	Means		
Log(Total individual income (in 1990 US dollars))	-0.124	0.088	-0.044	7.439		
Education(primary completed)						
Education(some secondary)						
Education (secondary or preparatory completed)						
Education (some higher education)						
Age (70-74)	-0.047	0.390	-0.017	0.275		
Age (75-79)	-0.498*	0.014	-0.187	0.142		
Age (80+)	-0.998*	0.155	-0.377	0.166		
Covered by social security						
Year dummy:1989	0.042*	0.015	0.015	0.519		
Constant	1.653*	0.575				

Table 26 (Urban) Probit of Probability of Living with Others: pooled 0: living alone 1: living with others

	Biprobit Urban males	i	Biprobit Selection		
	N=744			N=955	
		Standard			Standard
Variables	Coefficients	Errors	Variables	Coefficients	Errors
Log(Total individual income (in 1990 US dollars))	-0.026	0.029	Years of schooling	0.097*	0.014
Education(primary completed)	0.371*	0.103	Mean age of the household	0.023*	0.002
Education(some secondary)	0.117	0.165	Mean labor market status	0.285*	0.136
Education (secondary or preparatory completed)	0.160	0.260	of the household		
Education (some higher education)	0.128	0.188	Constant	-0.865*	0.178
Age (70-74)	-0.104	0.059			
Age (75-79)	-0.182*	0.074			
Age (80+)	-0.419*	0.110			
Covered by social security	0.014	0.060			
Year dummy:1989	0.457*	0.057			
Constant	0.028	0.203			

		_		le 27		
		The	projection of I	iving arrangements		
Mexico: 1984-1989		Rural fem	nales			
		Means	Means	-		
Variables	Coefficients	in 1984	in 1989	_		
Education(primary completed)	-0.341	0.069	0.041	Estimated probabilities of living with others		The explanatory power
Education(some secondary)	-2.233*	0.005	0.002	1984	0.827	of each independent variable
Education (secondary or preparatory completed)	-0.32	0.015	0.002	1989	0.835	for the change in the
Education (some higher education)						estimated probability
Age (70-74)	-0.052	0.31	0.279	Estimated probabilities at the mean values in	1984 except t	he following:
Age (75-79)	-0.254	0.153	0.171	Covered by social security	0.834	87.50%
Age (80+)	0.287	0.266	0.22	Age (75-79)	0.826	-12.50%
Covered by social security	0.271*	0.108	0.218	Education(some secondary)	0.829	25%
Year dummy:1995	-0.228*			•		
Constant	0.930*			_		
Mexico	: 1984-1989	Rural ma	les			
		Means	Means	-		
Variables	Coefficients	in 1984	in 1989	_		
Education(primary completed)	-0.399*	0.05	0.32	Estimated probabilities of living with others		
Education(some secondary)	0.185	0.01	0.011	1984	0.774	
Education (secondary or preparatory completed)	-0.270	0	0.005	1989	0.747	
Education (some higher education)	-1.303	0.005	0.004			
Age (70-74)	-0.205*	0.279	0.241	Estimated probabilities at the mean values in	1984 except t	he following:
Age (75-79)	-0.284*	0.214	0.216	Covered by social security	0.779	-18.50%
Age (80+)	0.044	0.22	0.194	Age (75-79)	0.774	0%
Covered by social security	0.225	0.134	0.198	Education(primary completed)	0.740	126%
Year dummy:1995	-0.150*			Education(some secondary)	0.773	3.70%
Constant	0.872*					

Table 28

Probit of Probability of Living with Others: pooled

0: living alone 1: living with others

Uruguay : 1981,1995

	all females		N=	7134	all females		N=	7134
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.073*	0.021	0.028	0.312	0.056	0.038	0.021	0.312
Education(some secondary)	-0.031	0.048	-0.012	0.083	-0.060*	0.013	-0.023	0.083
Education (secondary or preparatory completed)	-0.028	0.103	-0.011	0.015	-0.011	0.063	-0.004	0.015
Education (some higher education)	0.027	0.066	0.010	0.053	0.091	0.057	0.034	0.053
Age (70-74)	0.070*	0.007	0.027	0.271	-0.015	0.009	-0.006	0.271
Age (75-79)	0.111*	0.062	0.042	0.196	-0.056	0.066	-0.021	0.196
Age (80+)	0.430*	0.110	0.158	0.240	0.180	0.125	0.068	0.240
Log(Social security income (in 1990 US dollars))	-0.165*	0.005	-0.064	7.126	-0.271*	0.045	-0.104	7.126
Log(Social security income)*single					0.032*	0.003	0.012	0.753
Log(Social security income)*widow					0.038*	0.005	0.015	4.187
Log(Social security income)*married					-0.092*	0.005	-0.035	1.822
Year dummy:1995*single					0.030*	0.003	0.012	0.059
Year dummy:1995*widow					-0.105*	0.010	-0.040	0.343
Year dummy:1995*married					-0.015	0.055	-0.006	0.165
Year dummy:1995	-0.203*	0.010	-0.078	0.601	-0.096*	0.035	-0.037	0.601
constant	1.404*	0.069			2.239*	0.317		

	**Note: 2. tes	t			**Note:3. test			
	all females		N=	7134	all females		N=	7134
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.057	0.037	0.022	0.312	0.070*	0.020	0.027	0.312
Education(some secondary)	-0.057*	0.014	-0.022	0.083	-0.037	0.045	-0.014	0.083
Education (secondary or preparatory completed)	-0.005	0.070	-0.002	0.015	-0.00002	0.087	0.000	0.015
Education (some higher education)	0.090	0.058	0.034	0.053	0.058	0.030	0.022	0.053
Age (70-74)	-0.014	0.009	-0.006	0.271	0.024	0.050	0.009	0.271
Age (75-79)	-0.055	0.067	-0.021	0.196	0.019	0.149	0.007	0.196
Age (80+)	0.178	0.124	0.067	0.240	0.278	0.240	0.104	0.240
Log(Social security income (in 1990 US dollars))	-0.359*	0.038	-0.138	7.126	-0.234*	0.066	-0.090	7.126
Log(Social security income)*single	0.126*	0.005	0.049	0.753				
Log(Social security income)*widow	0.122*	0.002	0.047	4.187				
Log(Social security income)*married	0.092*	0.001	0.035	0.362				
Year dummy:1995*single					0.889	0.065	0.279	0.059
Year dummy:1995*widow					0.642	0.028	0.215	0.034
Year dummy:1995*married					0.791	0.082	0.286	0.343
Year dummy:1995	-0.157*	0.015	-0.060	0.601	-0.753	0.030	-0.277	0.601
constant	2.248*	0.306			1.945	0.554		

Note:1. When the same regression is analyzed by education groups, the coefficients do not change very much. (by 0.02 at the most for social security income) *These tests reject the similarity of the coefficients of social security income by marital status.

Note: 2. test The statistic is 533.73. Note: 3. test

Prob > chi2 = 0.0000

The statistic is 92.35.

Prob > chi2 = 0.0000

	Probit of Pro		le 29 .iving with Oth	are: pooled				
		,	living with oth	•				
Uruguay: 1981,1995	0. 110	ing alone 1.	iiviiig with oth	CIS				
oragaay. 1001,1000								
	all females		N=7,134		all females		N=7,134	
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)					0.057	0.034	0.022	0.312
Education(some secondary)					-0.059*	0.017	-0.023	0.083
Education (secondary or preparatory completed)					-0.016	0.073	-0.006	0.015
Education (some higher education)					0.081	0.056	0.031	0.053
Age (70-74)	0.071*	0.003	0.027	0.271	-0.013*	0.006	-0.005	0.271
Age (75-79)	0.110	0.058	0.042	0.196	-0.054	0.064	-0.021	0.196
Age (80+)	0.429*	0.103	0.158	0.240	0.185	0.123	0.070	0.240
a dummy for married people	0.405+	0.040	0.001	7.000	-0.854*	0.005	-0.330	0.261
Log(Social security income (in 1990 US dollars))	-0.165*	0.018	-0.064	7.283	-0.258*	0.038	-0.099	7.126
Log(Social security income)*single								
Log(Social security income)*widow Log(Social security income)*married								
Log(Total individual income (in 1990 US dollars))								
Year dummy:1995*single								
Year dummy:1995*widow								
Year dummy:1995*married								
Year dummy:1995	-0.203*	0.009	-0.078	0.601	-0.159*	0.014	-0.060	0.601
constant	1.424*	0.161	0.070	0.001	2.388*	0.299	0.000	0.001
constant		0.101			2.000	0.200		
	all females		N=7,134					
		Standard						
Variables	Coefficients	Errors	Derivatives	Means				
Education(primary completed)								
Education(some secondary)								
Education (secondary or preparatory completed)								
Education (some higher education)	0.04 **	0.00:	0.00-	0.074				
Age (70-74)	-0.014*	0.004	-0.005	0.271				
Age (75-79)	-0.056	0.062	-0.022	0.196				
Age (80+)	0.185 -0.853*	0.119 0.003	0.070 -0.329	0.240 0.261				
a dummy for married people Log(Social security income (in 1990 US dollars))	-0.853** -0.252*	0.003	-0.329 -0.097	7.126				
Log(Social security income)*single	-0.252	0.040	-0.097	1.120				
Log(Social security income) single								
Log(Social security income) *married								
Log(Total individual income (in 1990 US dollars))								
Year dummy:1995*single								
Year dummy:1995*widow								
Year dummy:1995*married								
Year dummy:1995	-0.159*	0.014	-0.061	0.601				
constant	2.366*	0.377						

Table 30 Probit of Probability of Living with Others: pooled 0: living alone 1: living with others : 1981,1995

Uruguay

	all females		N=	7404	all females		N=	7404
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.103*	0.008	0.040	0.312				
Education(some secondary)	0.012	0.022	0.005	0.088				
Education (secondary or preparatory completed)	0.035	0.094	0.014	0.016				
Education (some higher education)	0.102	0.057	0.039	0.053				
Age (70-74)	0.064*	0.004	0.024	0.269	0.062*	0.001	0.024	0.269
Age (75-79)	0.116	0.071	0.044	0.193	0.111	0.068	0.042	0.193
Age (80+)	0.430*	0.109	0.159	0.235	0.425*	0.103	0.157	0.235
Dummy variable for married people								
Log(Social security income (in 1990 US dollars))								
Log(Social security income)*single								
Log(Social security income)*widow								
Log(Social security income)*married								
Log(Total individual income (in 1990 US dollars))	-0.221*	0.022	-0.085	7.240	-0.208*	0.031	-0.080	7.240
Year dummy:1995*single	-0.221	0.022	-0.003	7.240	-0.200	0.031	-0.000	7.240
Year dummy:1995*widow								
Year dummy:1995 widow Year dummy:1995*married								
Year dummy:1995 married	-0.182*	0.013	-0.070	0.601	-0.180*	0.012	-0.069	0.601
constant	-0.162* 1.795*	0.013	-0.070	0.601	-0.160 1.743*	0.012	-0.069	0.601
Constant	1.790	0.197			1.743	0.237		
	all females		N=	7134	all females		N=	7134
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	
Education(primary completed)					0.098*	0.023	0.037	0.312
Education(some secondary)					0.028	0.035	0.011	0.083
Education (secondary or preparatory completed)					0.066	0.083	0.025	0.015
Education (some higher education)					0.100*	0.045	0.038	0.053
Age (70-74)	0.064*	0.011	0.025	0.271	0.067*	0.015	0.025	0.271
Age (75-79)	0.095	0.066	0.036	0.196	0.102	0.069	0.039	0.196
Age (80+)	0.421*	0.104	0.155	0.240	0.429*	0.109	0.158	0.240
Dummy variable for married people								
Log(Social security income (in 1990 US dollars))	0.104*	0.003	0.040	7.126	0.103*	0.005	0.039	7.126
Log(Social security income)*single								
Log(Social security income)*widow								
Log(Social security income)*married								
Log(Total individual income (in 1990 US dollars))	-0.307*	0.033	-0.118	7.236	-0.320*	0.026	-0.123	7.236
Year dummy:1995*single	0.001	2.000	20	00	3.020	2.020	220	200
Year dummy:1995*widow								
Year dummy:1995*married								
Year dummy:1995	-0.182*	0.012	-0.070	0.601	-0.185*	0.013	-0.071	0.601
constant	-0.162 1.721*	0.012	-0.070	0.001	1.789*	0.013	-0.071	0.001
	1 /21*	0.249			1 /89*	0.189		

			ole 31					
			Living with Oth					
		ing alone 1	: living with ot	hers				
Uruguay	: 1981,1995							
	females: sin	gle	N=	736	females: sin	gle	N=	736
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	-0.051	0.034	-0.017	0.307				
Education(some secondary)	-0.222*	0.061	-0.079	0.099				
Education (secondary or preparatory completed)	-0.094	0.241	-0.033	0.022				
Education (some higher education)	0.327	0.300	0.102	0.087				
Age (70-74)	0.251*	0.018	0.082	0.254	0.230*	0.005	0.076	0.254
Age (75-79)	-0.011	0.195	-0.004	0.185	-0.037	0.180	-0.013	0.185
Age (80+)	0.311	0.177	0.101	0.261	0.299	0.174	0.097	0.261
Log(Social security income (in 1990 US dollars)) Log(Total individual income (in 1990 US dollars))	-0.354*	0.046	-0.120	7.115	-0.303*	0.096	-0.103	7.115
Year dummy:1995	-0.056	0.015	-0.019	0.555	-0.075*	0.032	-0.025	0.555
constant	2.988*	0.409	-0.013	0.555	2.637*	0.752	-0.023	0.555
	females: sin	gle	N=	780	females: sin	gle	N=	780
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.025*	0.011	0.008	0.309				
Education(some secondary)	-0.069	0.120	-0.023	0.105				
Education (secondary or preparatory completed)	0.064	0.229	0.021	0.021				
Education (some higher education)	0.548*	0.218	0.156	0.086				
Age (70-74)	0.296*	0.089	0.095	0.255	0.264*	0.084	0.085	0.255
Age (75-79)	-0.069	0.231	-0.023	0.182	-0.114	0.228	-0.039	0.182
Age (80+)	0.313	0.247	0.100	0.255	0.284	0.260	0.092	0.255
Log(Social security income (in 1990 US dollars))								
Log(Total individual income (in 1990 US dollars))	-0.462*	0.067	-0.155	7.245	-0.368*	0.085	-0.124	7.245
Year dummy:1995	0.014	0.024	0.005	0.552	-0.002	0.040	-0.001	0.552
	3.733*	0.605			3.131*	0.728		

Table 32
Probit of Probability of Living with Others: pooled
0: living alone 1: living with others
: 1981,1995

Uruguay

	females: div	orced	N=	376	females: div	orced	N=	376
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.271*	0.119	0.102	0.292				
Education(some secondary)	-0.163	0.509	-0.063	0.116				
Education (secondary or preparatory completed)	-0.027	0.046	-0.011	0.030				
Education (some higher education)	0.008	0.017	0.003	0.067				
Age (70-74)	0.098*	0.001	0.037	0.303	0.089*	0.016	0.034	0.303
Age (75-79)	0.117	0.085	0.044	0.198	0.121	0.064	0.046	0.198
Age (80+)	0.176	0.271	0.066	0.128	0.202	0.207	0.076	0.128
Log(Social security income (in 1990 US dollars))	-0.337*	0.128	-0.129	7.113	-0.343*	0.097	-0.132	7.113
Log(Total individual income (in 1990 US dollars))								
Year dummy:1995	-0.097*	0.011	-0.037	0.668	-0.070*	0.004	-0.027	0.668
constant	2.601*	0.868			2.688*	0.669		
	females: div	orced	N=	423	females: div	orced	N=	423
		Standard				Standard		

	females: div	orced	N=	423	females: div	orced	N=	423
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.342*	0.062	0.128	0.290				
Education(some secondary)	0.065	0.481	0.025	0.139				
Education (secondary or preparatory completed)	0.217*	0.023	0.081	0.034				
Education (some higher education)	0.066	0.059	0.025	0.068				
Age (70-74)	-0.0004	0.015	-0.0002	0.292	-0.006	0.032	-0.002	0.292
Age (75-79)	0.059	0.158	0.023	0.189	0.076	0.163	0.029	0.189
Age (80+)	0.080	0.167	0.030	0.116	0.083	0.085	0.032	0.116
Log(Social security income (in 1990 US dollars))								
Log(Total individual income (in 1990 US dollars))	-0.434*	0.172	-0.167	7.306	-0.423*	0.115	-0.163	7.306
Year dummy:1995	0.002	0.009	0.001	0.678	0.049*	0.022	0.019	0.678
constant	3 300*	1 215			3 301*	0.860		

		Tab	le 33					
	Probit of Pr	obability of L	iving with Oth	ers: pooled				
	0: li	ving alone 1:	living with oth	ners				
Uruguay	: 1981,1995							
	females: wid	low	N=	4124	females: wid	low	N=	4124
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.127*	0.017	0.044	0.321				
Education(some secondary)	0.080*	0.034	0.028	0.080				
Education (secondary or preparatory completed)	0.065	0.195	0.022	0.012				
Education (some higher education)	0.032	0.024	0.011	0.045				
Age (70-74)	0.040	0.033	0.014	0.255	0.038	0.038	0.013	0.255
Age (75-79)	0.030*	0.014	0.010	0.218	0.024	0.016	0.008	0.218
Age (80+)	0.288*	0.090	0.098	0.312	0.280*	0.083	0.095	0.312
Log(Social security income (in 1990 US dollars))	-0.276*	0.028	-0.097	7.189	-0.268*	0.032	-0.094	7.189
Log(Total individual income (in 1990 US dollars))								
Year dummy:1995	-0.210*	0.017	-0.072	0.589	-0.205*	0.015	-0.071	0.589
constant	2.467*	0.210			2.457*	0.239		
	females: wid	low	N=	4196	females: wid	low	N=	4196
		Standard				Standard		
Variables	Coefficients	Errors	Derivatives	Means	Coefficients	Errors	Derivatives	Means
Education(primary completed)	0.194*	0.027	0.066	0.322				
Education(some secondary)	0.236*	0.012	0.077	0.081				
Education (secondary or preparatory completed)	0.259*	0.131	0.084	0.012				
Education (some higher education)	0.218*	0.082	0.072	0.045				
Age (70-74)	0.042	0.029	0.015	0.254	0.036	0.037	0.013	0.254
Age (75-79)	0.029	0.037	0.010	0.218	0.016	0.038	0.006	0.218
Age (80+)	0.266	0.093	0.090	0.310	0.249*	0.086	0.084	0.310
Log(Social security income (in 1990 US dollars))			-0.154	7.329				
Log(Total individual income (in 1990 US dollars))	-0.442*	0.098			-0.396*	0.083	-0.138	7.329
Year dummy:1995	-0.159*	0.032	-0.055	0.586	-0.154*	0.029	-0.053	0.586
constant	3.663*	0.721			3.418*	0.624		

 Table 34

 The projection of living arrangements

Uruguay	: 1981-1995	Females:	einalo			
Oruguay	. 1901-1993	Means	Means			
Variables	Coefficients	in 1981	in 1995			
Education(primary completed)	-0.051	0.337	0.292	Estimated probabilities of living with others		The explanatory power
	-0.222*	0.337	0.292	1981	0.736	
Education(some secondary)						of each independent variable
Education (secondary or preparatory completed)	-0.094	0.004	0.036	1995	0.716	for the change in the
Education (some higher education)	0.327	0.081	0.086	F		estimated probability
Age (70-74)	0.251*	0.252	0.258	Estimated probabilities at the mean values in 1984 excep		4000/
Age (75-79)	-0.011	0.175	0.191	Social security income (in 1990 US dollars)	0.716	100%
Age (80+)	0.311	0.240	0.265	Age (75-79)	0.736	0%
Log(Social security income (in 1990 US dollars))	-0.354*	7.022	7.188	Education(some secondary)	0.732	20%
Log(Total individual income (in 1990 US dollars))				Education (secondary or preparatory completed)	0.735	5%
Year dummy:1995	-0.056			Doubled social security income	0.633	
constant	2.988*			\$1000 more social security income	0.648	
				\$500 more social security income	0.680	
Uruguay	: 1981-1995	Females:	divorced			
		Means	Means			
Variables	Coefficients	in 1981	in 1995			
Education(primary completed)	0.271*	0.245	0.326	Estimated probabilities of living with others		
Education(some secondary)	-0.163	0.085	0.169	1981	0.641	
Education (secondary or preparatory completed)	-0.027	0.021	0.040	1995	0.625	
Education (some higher education)	0.008	0.041	0.050			
Age (70-74)	0.098*	0.300	0.280	Estimated probabilities at the mean values in 1984 except	t the following:	
Age (75-79)	0.117	0.234	0.163	Social security income (in 1990 US dollars)	0.622	119%
Age (80+)	0.176	0.085	0.131	Age (75-79)	0.638	19%
Log(Social security income (in 1990 US dollars))	-0.337*	7.012	7.161	Education(some secondary)	0.636	31%
Log(Total individual income (in 1990 US dollars))				Education (secondary or preparatory completed)	0.641	0%
Year dummy:1995	-0.097*			Doubled social security income	0.553	0 / 0
constant	2.601*			\$1000 more social security income	0.559	
Constant	2.001			\$500 more social security income	0.595	
Uruguay	: 1981-1995	Females:	widow			
Oraguay	. 1001-1000	Means	Means			
Variables	Coefficients	in 1981	in 1995	Estimated probabilities of living with others		
Education(primary completed)	0.127*	0.318	0.329	1981	0.746	
Education(some secondary)	0.080*	0.066	0.09	1995	0.733	
Education (secondary or preparatory completed)	0.065	0.007	0.016		0.013	
Education (some higher education)	0.032	0.032	0.083	Estimated probabilities at the mean values in 1984 except	t the following:	
Age (70-74)	0.040	0.274	0.241	Social security income (in 1990 US dollars)	0.725	162%
Age (75-79)	0.030*	0.219	0.213	Age (75-79)	0.746	0%
Age (80+)	0.288*	0.267	0.342	Education(some secondary)	0.747	-0.08%
Log(Social security income (in 1990 US dollars))	-0.276*	7.050	7.285	Education (secondary or preparatory completed)	0.746	0%
Log(Total individual income (in 1990 US dollars))				Doubled social security income	0.661	
Year dummy:1995	-0.210*			\$1000 more social security income	0.677	
constant	2.467*			\$500 more social security income	0.700	
oonstant	2.401			4000 more addian accurity income	0.700	

Table A.1: The ratio of social security income to total income: Uruguay 1981

	tota	al			fema	les			male	es	
age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.
65-69	0.92	0.21	1031	65-69	0.95	0.17	583	65-69	0.88	0.25	448
70-74	0.93	0.19	962	70-74	0.95	0.17	532	70-74	0.91	0.21	430
75-79	0.95	0.17	647	75-79	0.96	0.15	371	75-79	0.93	0.19	276
80-	0.96	0.15	632	80-	0.97	0.13	403	80-	0.94	0.18	229
Total	0.94	0.19	3272	Total	0.95	0.16	1889	Total	0.91	0.22	1383
					females:	single			females: ı	married	
				age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.
				65-69	0.91	0.22	75	65-69	0.98	0.09	211
				70-74	0.94	0.18	56	70-74	1.00	0.03	140
				75-79	0.95	0.18	38	75-79	0.99	0.04	63
				80-	0.98	0.11	54	80-	1.00	0.00	37
				Total	0.94	0.18	223	Total	0.99	0.07	451
					females: c	livorced			females:	widow	
				age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs
				65-69	0.98	0.08	28	65-69	0.93	0.20	268
				70-74	0.98	0.08	25	70-74	0.93	0.21	311
				75-79	0.96	0.13	19	75-79	0.95	0.16	251
				80-	1.00	0.00	8	80-	0.96	0.15	304
				Total	0.98	0.09	80	Total	0.94	0.18	113

	tota	l			fema	les			mal	es	
age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.
65-69	0.88	0.23	2549	65-69	0.90	0.22	1481	65-69	0.85	0.23	1068
70-74	0.91	0.20	2333	70-74	0.92	0.19	1384	70-74	0.89	0.21	949
75-79	0.93	0.17	1690	75-79	0.94	0.16	1026	75-79	0.92	0.17	664
80-	0.92	0.19	2034	80-	0.92	0.19	1354	80-	0.92	0.19	680
Total	0.91	0.20	8606	Total	0.92	0.19	5245	Total	0.89	0.21	3361
	females: single females: married					females: single females: n					
				age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.
				65-69	0.89	0.23	139	65-69	0.96	0.15	633
				70-74	0.92	0.19	132	70-74	0.98	0.09	441
				75-79	0.94	0.18	101	75-79	0.99	0.05	228
				80-	0.96	0.13	141	80-	0.99	0.04	144
				Total	0.93	0.19	513	Total	0.97	0.12	1446
					females: c	livorced			females:	widow	
				age group	Ratio	std.dev.	obs.	age group	Ratio	std.dev.	obs.
				65-69	0.80	0.27	113	65-69	0.86	0.25	596
				70-74	0.85	0.27	88	70-74	0.90	0.21	723
				75-79	0.90	0.23	53	75-79	0.92	0.17	644
				80-	0.93	0.13	42	80-	0.91	0.21	1027
				Total	0.85	0.25	296	Total	0.90	0.21	2990

Table A.2: The ratio of social security income to total income: Uruguay 1995

Table A.3: Percentage of elderly living alone: Latin America														
country	Argentina	Bolivia	Brazil	Chile	Colombia	Costa Rica	Ecuador	Guatemala	Honduras	Mexico	Nicaragua	Panama	Peru	Paraguay
year	1996	1997	1997	1998	1998	1998	1998	1998	1998	1996	1998	1999	1996	1998
total	0.51	0.41	0.33	0.30	0.19	0.32	0.24	0.21	0.15	0.26	0.12	0.27	0.24	0.18
females	0.48	0.39	0.31	0.29	0.17	0.28	0.23	0.17	0.13	0.24	0.10	0.22	0.23	0.17
males	0.54	0.42	0.36	0.33	0.22	0.35	0.25	0.24	0.16	0.28	0.14	0.33	0.26	0.19
rural		0.51	0.33	0.27	0.25	0.31	0.25	0.19	0.15	0.27	0.12	0.32	0.33	0.18
urban		0.31	0.33	0.31	0.15	0.32	0.23	0.22	0.14	0.25	0.12	0.25	0.19	0.19
Table A.4: Probit of proba country year	Argentina 1996	Bolivia 1997	Brazil 1997	Chile 1998	Colombia 1998	Costa Rica 1998	Ecuador 1998	Guatemala 1998	Honduras 1998	Mexico 1996	Nicaragua 1998	Panama 1999	Peru 1996	Paraguay 1998
missing income (%)	19.62	42.33	11.48	12.91	44.58	28.38	44.25	25.53	41.73	42.31	73.9	27.49	28.05	37.39
coefficients of total individual income	-0.086*	0.010	-0.063*	-0.121*	0.058*	-0.061	0.058	0.046	0.006	0.056*	0.013	0.026	0.005	-0.016
standard errors Obs. in probit	0.024 6889	0.036 1200	0.014 17161	0.014 13159	0.022 4105	0.038 1833	0.032 799	0.034 1470	0.039 824	0.028 1723	0.103 251	0.032 2024	0.025 2751	0.058 704

Note: 1. Persons above 65 with non-missing income are considered for the probit analysis.

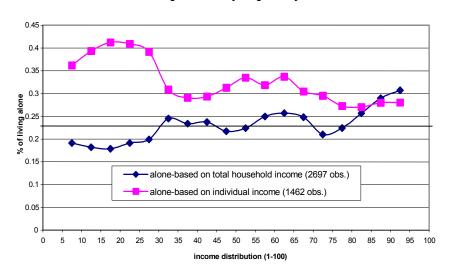
Note: 2. Gender, region, education dummies (primary completed, some secondary, secondary completed, and some higher education), age group dummies (70-74, 75-79, 80+), and log (total individual income) are included as explanatory variables.

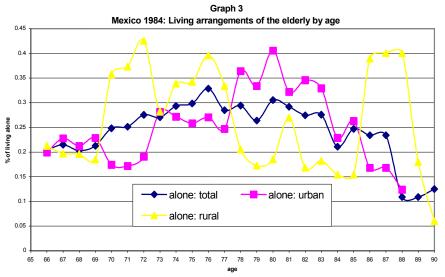
Note: 3. In the Venezuelan survey, urban areas are not distinguishable from rural areas. Argentina and Uruguay are urban only.

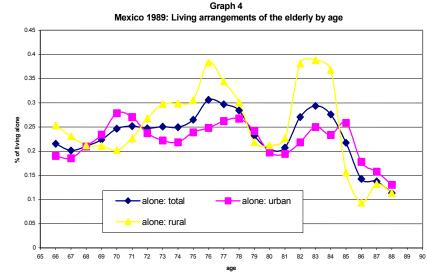
Note: 4. In the survey of Nicaragua, income from all jobs is utilized instead of income from all sources.

Graph 1
Mexico 1984: Percentage of the elderly living alone by income distribution

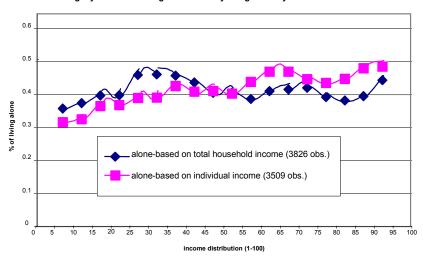
Graph 2
Mexico 1989: Percentage of the elderly living alone by income distribution



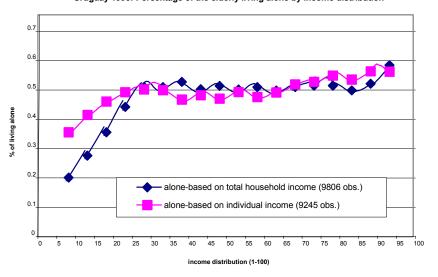




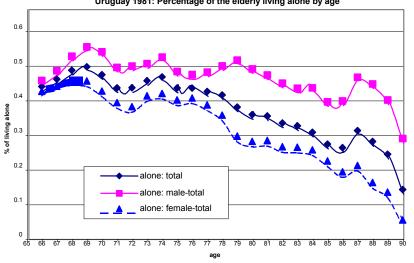
Graph 5
Uruguay 1981: Percentage of the elderly living alone by income distribution



Graph 6
Uruguay 1995: Percentage of the elderly living alone by income distribution



Graph 7
Uruguay 1981: Percentage of the elderly living alone by age



Graph 8
Uruguay 1995: Percentage of the elderly living alone by age

